

# Exponential inequalities for martingales and asymptotic properties of the free energy of directed polymers in random environment

Quansheng Liu\* and Frédérique Watbled†

LMAM, Université de Bretagne Sud,  
Campus de Tohannic, BP 573, 56017 Vannes, France  
Université Européenne de Bretagne

November 2008

## Abstract

We first obtain exponential inequalities for martingales. Let  $(X_k)(1 \leq k \leq n)$  be a sequence of martingale differences relative to a filtration  $(\mathcal{F}_k)$ , and set  $S_n = X_1 + \dots + X_n$ . We prove that if for some  $\delta > 0, Q \geq 1, K > 0$  and all  $k$ , a.s.  $\mathbb{E}[e^{\delta|X_k|^Q} | \mathcal{F}_{k-1}] \leq K$ , then for some constant  $c > 0$  (depending only on  $\delta, Q$  and  $K$ ) and all  $x > 0$ ,  $P[|S_n| > nx] \leq 2e^{-nc(x)}$ , where  $c(x) = cx^2$  if  $x \in ]0, 1]$ , and  $c(x) = cx^Q$  if  $x > 1$ ; the converse also holds if  $(X_i)$  are independent and identically distributed. It extends Bernstein's inequality for  $Q = 1$ , and Hoeffding's inequality for  $Q = 2$ . We then apply the preceding result to establish exponential concentration inequalities for the free energy of directed polymers in random environment, show its rate of convergence (in probability, almost surely, and in  $L^p$ ), and give it an expression in terms of free energies of some multiplicative cascades, which improves an inequality of Comets and Vargas (2006, [13]) to an equality.

**Key words.** Martingale differences, super-martingales, large deviation inequality, exponential inequality, Bernstein's inequality, Hoeffding-Azuma's inequality, directed polymers, random environment, concentration inequality, free energy, convergence rate, multiplicative cascades.

**2000 AMS Subject Classifications.** Primary 60G42, 60K35; secondary 60K37, 60G50, 60F10, 82D30

## 1 Introduction and main results

Our work was initially motivated by the study of the free energy of a directed polymer in a random environment. Comets and Vargas (2006, [13]) proved that the free energy (at  $\infty$ ) is bounded by the infimum of those of some generalized multiplicative cascades, and that the equality holds if the environment is bounded or gaussian. The essential point in their proof for the equality is an exponential concentration inequality for the free energy (at time  $n$ ), which was not known for a general environment. Using a large deviation inequality of Lesigne and Volny (2001, [27]) on martingales, Comets, Shiga and Yoshida (2003, [11]) did obtain a concentration inequality for the free energy; but their bound is larger than the exponential one, and is not sharper enough to imply the equality mentioned above. Another non satisfactory point of their inequality is that it cannot be used to prove rigorous results on the rate of convergence, for the almost sure (a.s.) or  $L^p$  convergence of the free energies.

The objective of the present paper is to establish exponential large deviation inequalities, and to use them to show exponential concentration inequalities for the free energy of a polymer in general random environment, its rate of convergence, and an expression of its limit value in terms of those of some multiplicative cascades.

---

\*Email: Quansheng.Liu@univ-ubs.fr

†Email: Frederique.Watbled@univ-ubs.fr

Large deviation inequalities are very powerful tools in probability theory, and have been studied by many authors: see e.g. the classical works of Bernstein (1924, [3]), Cramér (1938, [15]), Hoeffding (1963, [22]), Azuma (1967, [1]), Chernoff (1981, [8]), the books of Chow and Teicher (1978, [9]), and Petrov (1995, [31]), and the recent papers by de la Peña, (1999, [16]), Lesigne and Volný (2001, [27]), Bentkus (2004, [2]), and Chung and Lu (2006, [10]). See also Ledoux (1999, [26]) and Wang (2005, [33]) for related concentration inequalities and general functional inequalities.

Let  $(\Omega, \mathcal{F}, P)$  be a probability space, and let  $\mathcal{F}_0 = \{\emptyset, \Omega\} \subset \mathcal{F}_1 \subset \dots \subset \mathcal{F}_n$  be an increasing sequence of sub- $\sigma$ -fields of  $\mathcal{F}$ . Let  $X_1, \dots, X_n$  be a sequence of real-valued martingale differences defined on  $(\Omega, \mathcal{F}, P)$ , adapted to the filtration  $(\mathcal{F}_k)$ : that is, for each  $1 \leq k \leq n$ ,  $X_k$  is  $\mathcal{F}_k$  measurable, and  $\mathbb{E}[X_k | \mathcal{F}_{k-1}] = 0$ . Set

$$S_n = X_1 + \dots + X_n. \quad (1.1)$$

We are interested in exponential large deviation inequalities of the form

$$P[|S_n| > nx] = O(e^{-c(x)n}), \quad (1.2)$$

where  $x > 0$  and  $c(x) > 0$ . When  $(X_i)$  are independent and identically distributed (iid) with mean  $\mathbb{E}X_i = 0$ , it is known [see Petrov (1995, [31] p.137)] that (1.2) holds for all  $x > 0$  and some  $c(x) > 0$  if and only if for some  $\delta > 0$ ,

$$\mathbb{E}e^{\delta|X_1|} < \infty. \quad (1.3)$$

For a sequence of martingale differences, Lesigne and Volný (2001, [27]) proved that if for some constant  $K > 0$  and all  $k = 1, \dots, n$ ,

$$\mathbb{E}e^{|X_k|} \leq K, \quad (1.4)$$

then for any  $x > 0$ ,

$$P\left[\frac{S_n}{n} > x\right] = O(e^{-\frac{1}{4}x^{2/3}n^{1/3}}), \quad (1.5)$$

and that this is the best possible inequality that we can have under the condition (1.4), even in the class of stationary and ergodic sequences of martingale differences, in the sense that there exist such sequences of martingale differences  $(X_i)$  satisfying (1.4) for some  $K > 0$ , but

$$P\left[\frac{S_n}{n} > 1\right] > e^{-cn^{1/3}} \quad (1.6)$$

for some constant  $c > 0$  and infinitely many  $n$ . It is therefore interesting to know what is the good condition to have the exponential inequality (1.2) in the martingale case. It turns out that (1.2) still holds if we replace the expectation in (1.4) by the conditional one given  $\mathcal{F}_{k-1}$ . In fact we shall prove the following much sharper result. It is a consequence of Theorems 2.1, 3.1, and 3.2.

**THEOREM 1.1** *Let  $(X_k)$  be a  $\{\mathcal{F}_k\}$ -adapted sequence of martingale differences. Assume that for some constants  $Q \geq 1$ ,  $\delta > 0$ ,  $K > 0$  and all  $k \in \{1, \dots, n\}$ , almost surely*

$$\mathbb{E}[e^{\delta|X_k|^Q} | \mathcal{F}_{k-1}] \leq K. \quad (1.7)$$

*Then there exists a constant  $c > 0$  depending only on  $Q$ ,  $\delta$  and  $K$ , such that for all  $x > 0$ ,*

$$P\left[\pm \frac{S_n}{n} > x\right] \leq \begin{cases} e^{-ncx^2} & \text{if } x \in ]0, 1], \\ e^{-ncx^Q} & \text{if } x \in ]1, \infty[. \end{cases} \quad (1.8)$$

*The converse also holds in the iid case: if  $X_k$  are iid and if  $P[\pm \frac{S_n}{n} > x] \leq e^{-ncx^Q}$  holds for some  $n \geq 1$ ,  $Q \geq 1$ ,  $c > 0$ ,  $x_1 > 0$  and all  $x \geq x_1$ , then for all  $\delta \in ]0, c[$ , there exists  $K = K(\delta, Q, c, x_1) > 0$  such that*

$$\mathbb{E}[e^{\delta|X_1|^Q}] \leq K.$$

By the result of Lesigne and Volný ([27]) cited above, the conditional exponential moment condition (1.7) cannot be relaxed to the non conditional one.

When  $(X_k)$  are iid with  $\mathbb{E}[X_k] = 0$ , Bernstein's inequality states (cf. [31], page 57) that if  $\sigma^2 = \mathbb{E}[X_k^2] < \infty$  and

$$|\mathbb{E}X_k^m| \leq \frac{1}{2}m!\sigma^2 H^{m-2} \quad (1.9)$$

for some  $H > 0$  and all  $m = 2, 3, \dots$ , then

$$P\left[\pm \frac{S_n}{n} > x\right] \leq \begin{cases} e^{-nc_0x^2} & \text{if } x \in ]0, x_0], \\ e^{-nc_1x} & \text{if } x \in ]x_0, \infty[, \end{cases} \quad (1.10)$$

where  $c_0 = \frac{1}{4\sigma^2}$ ,  $c_1 = \frac{1}{4H}$ ,  $x_0 = \frac{\sigma^2}{H}$ . Notice that (in the iid case) Bernstein's condition (1.9) is equivalent to Cramer's condition that  $\exists \delta > 0$  such that

$$\mathbb{E}[e^{\delta|X_k|}] < \infty. \quad (1.11)$$

In applications we find more convenient to use Cramer's condition. Taking  $Q = 1$  in Theorem 1.1, we obtain the following Bernstein-type inequality.

**COROLLARY 1.2** (*A Bernstein-type inequality*) *Assume that  $(X_k)$  are iid with  $\mathbb{E}[X_k] = 0$ ,  $1 \leq k \leq n$ . If (1.11) holds for some  $\delta > 0$ , then for some  $c = c(\delta) > 0$ ,*

$$P\left[\pm \frac{S_n}{n} > x\right] \leq \begin{cases} e^{-ncx^2} & \text{if } x \in ]0, 1], \\ e^{-ncx} & \text{if } x \in ]1, \infty[. \end{cases} \quad (1.12)$$

*Conversely, if for some  $n \geq 1$ ,  $c > 0$ ,  $x_0 > 0$  and all  $x > x_0$ ,  $P\left[\pm \frac{S_n}{n} > x\right] \leq e^{-ncx}$ , then (1.11) holds for each  $\delta \in ]0, c[$ .*

When  $Q = 2$ , Theorem 1.1 extends the following well-known Hoeffding's inequality<sup>1</sup>: if  $(X_k)$  is a sequence of martingale differences with  $|X_k| \leq a$  a.s. for some constant  $a \in ]0, \infty[$ , then for all  $n \geq 1$  and all  $x > 0$ ,

$$P\left[\pm \frac{S_n}{n} > x\right] \leq e^{-ncx^2}, \quad (1.13)$$

where  $c = 1/(2a^2)$ . In fact, by our result for  $Q = 2$ , we obtain:

**COROLLARY 1.3** (*Extension of Hoeffding's inequality*) *When  $(X_k)$  are iid, then there is a constant  $c > 0$  such that (1.13) holds for all  $n \geq 1$  and all  $x > 0$ , if and only if for some  $\delta > 0$ ,*

$$\mathbb{E}e^{\delta X_1^2} < \infty. \quad (1.14)$$

*Moreover, if (1.13) holds for some  $n \geq 1$  and all  $x > 0$ , with some constant  $c = c_1$ , then it holds for all  $n \geq 1$  and  $x > 0$ , with some constant  $c = c_2$  depending only on  $c_1$ .*

So our result is a complete extension of Hoeffding's inequality even in the iid case.

We then apply the preceding results to directed polymers in random environment that we describe as follows. Let  $(\omega_n)_{n \in \mathbb{N}}$  be the simple random walk on  $\mathbb{Z}^d$  starting at 0, defined on a probability space  $(\Omega, \mathcal{F}, P)$ . Let  $(\eta(n, x))_{(n, x) \in \mathbb{N} \times \mathbb{Z}^d}$  be a sequence of i.i.d. real random variables defined on another probability space  $(E, \mathcal{E}, \mathbb{Q})$  (we use the letter E to refer the Environment). For real  $\beta$  (the inverse of temperature), define

$$\lambda(\beta) = \ln \mathbb{Q}[e^{\beta \eta(0, 0)}]. \quad (1.15)$$

(If  $\mu$  is a measure and  $f$  is a function, we write  $\mu(f)$  or  $\mu[f]$  for the integral of  $f$  with respect to  $\mu$ .) We fix  $\beta > 0$ , and only suppose that

$$\mathbb{Q}[e^{\beta |\eta(0, 0)|}] < \infty \quad (1.16)$$

(we do not suppose that it holds for all  $\beta > 0$ ). Of course this condition is equivalent to  $\lambda(\pm \beta) < \infty$ . We are interested in the normalized partition function

$$W_n(\beta) = P\left[\exp\left(\beta \sum_{j=1}^n \eta(j, \omega_j) - n\lambda(\beta)\right)\right], \quad (1.17)$$

---

<sup>1</sup>The inequality (1.13) is often called Hoeffding's inequality when  $(X_k)$  are iid, and Azuma's inequality when  $(X_k)$  are martingale differences. This is rather strange, as it was Hoeffding (1963) who first obtained it for martingales, although he mainly treated the iid case, and only mentioned the martingale case as a remark [see [22], p.18]. To respect the history, we call it Hoeffding's inequality, although Azuma (1967, [1]) refound it four years later. We think that what happened would be that, the first author who called it Azuma's inequality did not know the existence of the remark of Hoeffding, the second followed the first without verification, and so on.

and the free energy  $\frac{1}{n} \ln W_n(\beta)$ .

This model first appeared in physics literature [see Huse and Henley (1985, [23])] to modelize the phase boundary of Ising model subject to random impurities; the first mathematical study was undertaken by Imbrie and Spencer (1988, [24]) and Bolthausen (1989, [4]). For recent results, see e.g. Carmona and Hu (2004, [7]), Carmona, Guerra, Hu and Méjane (2006, [5]), Comets, Shiga and Yoshida (2004, [12]), and Comets and Yoshida (2006, [14]).

Assuming  $Q[e^{\beta|\eta(0,0)|}] < \infty$  for all  $\beta > 0$ , Comets, Shiga and Yoshida ([11]) proved that  $\forall x > 0$ , there exists  $n_0 \in \mathbb{N}^*$  such that for any  $n \geq n_0$ ,

$$\mathbb{Q} \left[ \left| \frac{1}{n} \ln W_n(\beta) - \frac{1}{n} \mathbb{Q}[\ln W_n(\beta)] \right| > x \right] \leq \exp \left( -\frac{n^{\frac{1}{3}} x^{\frac{2}{3}}}{4} \right). \quad (1.18)$$

In fact, in their proof of (1.18), they used the condition that  $Q[e^{3\beta|\eta(0,0)|}] < \infty$ , due to the application of their Lemma 3.1 (p.711).

We first improve this result to an exponential inequality under the weaker condition that  $Q[e^{\beta|\eta(0,0)|}] < \infty$  for the fixed  $\beta$ .

**THEOREM 1.4 (Exponential concentration inequality for the free energy)** *Let  $\beta > 0$  be fixed such that  $Q[e^{\beta|\eta(0,0)|}] < \infty$ . If for some  $Q \geq 1$  and  $R > 0$ ,*

$$\mathbb{Q}[e^{R|\eta(0,0)|^Q}] < +\infty, \quad (1.19)$$

then

$$\mathbb{Q} \left[ \frac{1}{n} |\ln W_n(\beta) - \mathbb{Q}[\ln W_n(\beta)]| > x \right] \leq \begin{cases} 2e^{-ncx^2} & \text{if } 0 \leq x \leq 1, \\ 2e^{-ncx^Q} & \text{if } x > 1, \end{cases} \quad (1.20)$$

where  $c > 0$  is a constant depending only on  $Q, R$ , and the law of  $\eta(0,0)$ .

Notice that the condition (1.19) holds automatically for  $Q = 1$  and  $R = \beta$ , so that (1.20) holds for  $Q = 1$  under the only hypothesis  $Q[e^{\beta|\eta(0,0)|}] < \infty$ ; when (1.19) holds for some  $Q > 1$  and  $R > 0$ , (1.20) gives a sharper bound for large values of  $x$ .

Theorem 1.4 is a consequence of Corollary 6.7. As shown in Carmona and Hu (2002, [6]) and Comets and Vargas (2006, [13]), when the environment is gaussian or bounded, the inequality can be obtained directly by a general concentration result on gaussian or bounded variables (see e.g. Ledoux (1999, [26])). But this method does not work for a general environment.

As applications we shall show the following properties about the free energy  $\frac{1}{n} \ln W_n(\beta)$ :

- (1) for some  $p_-(\beta) \leq 0$ ,  $\frac{1}{n} \ln W_n(\beta) \rightarrow p_-(\beta)$  in probability at an exponential rate (cf. Theorem 7.2);
- (2)  $\frac{1}{n} \ln W_n(\beta) \rightarrow p_-(\beta)$  a.s. and in  $L^p$ , for all  $p \geq 1$ , at a rate  $O\left(\sqrt{\frac{\ln n}{n}}\right)$  (cf. Theorem 7.5);
- (3)  $p_-(\beta)$  can be expressed in terms of some generalized multiplicative cascades (cf. Theorem 8.1).

Part (1) extends the same conclusion of Carmona and Hu (2002, [6]) for the gaussian environment case to a general environment case. The rate of a.s. convergence in part (2) improves the bound  $O(n^{-(\frac{1}{2}-\varepsilon)})$  ( $\varepsilon > 0$ ) of Carmona and Hu (2004, [7]) obtained for the gaussian environment case. Part (3) improves an inequality of Comets and Vargas (2006, [13]) to an equality.

The rest of the paper is organized as follows. In Section 2 we establish exponential inequalities for supermartingales, which extend Bernstein or Hoeffding's inequalities, according to  $\mathbb{E}[e^{\delta|X_i|} | \mathcal{F}_{i-1}] \leq K$  or  $\mathbb{E}[e^{\delta|X_i|^2} | \mathcal{F}_{i-1}] \leq K$ , respectively. For large values of  $x$ , sharper inequalities are proven in Section 3 under the condition that  $\mathbb{E}[e^{\delta|X_i|^Q} | \mathcal{F}_{i-1}] \leq K$  ( $Q > 1$ ). These results are extended in Section 4 to the more general case where  $\mathbb{E}[e^{\delta|X_i|^Q} | \mathcal{F}_{i-1}] \leq K_i$ . As applications, we show in Section 5 the rate of convergences, a.s. and in  $L^p$ . In the last 3 sections, we study the free energies of directed polymers in random environment, with the help of our results on martingales: we show exponential concentration inequalities for the free energies in Section 6, their convergence rates (in probability, a.s. and in  $L^p$ ) in Section 7, and, in Section 8, an expression of their limit value in terms of some generalized multiplicative cascades.

## 2 Exponential inequalities for supermartingales

In this section we give an extension of Bernstein and Hoeffding's inequalities to supermartingales with unbounded differences. Our results are sharp even in the iid case.

Let  $(X_i)_{1 \leq i \leq n}$  be a sequence of real-valued supermartingale differences defined on a probability space  $(\Omega, \mathcal{F}, P)$ , adapted to a filtration  $(\mathcal{F}_i)$ , with  $\mathcal{F}_0 = \{\emptyset, \Omega\}$ . This means that for each  $1 \leq i \leq n$ ,  $X_i$  is  $\mathcal{F}_i$ -measurable, and  $\mathbb{E}[X_i | \mathcal{F}_{i-1}] \leq 0$  a.s.. We are interested in the growth rate of the Laplace transform  $\mathbb{E}[e^{tS_n}]$ , and the convergence rate of the deviation probabilities  $P[\frac{S_n}{n} > x]$ .

**THEOREM 2.1** *Let  $(X_i)_{1 \leq i \leq n}$  be a finite sequence of supermartingale differences. If for some constant  $K > 0$  and all  $i \in [1, n]$ ,*

$$\mathbb{E}[e^{|X_i|} | \mathcal{F}_{i-1}] \leq K \quad \text{a.s.}, \quad (2.1)$$

*then:*

$$\mathbb{E}[e^{tS_n}] \leq \exp\left(\frac{nKt^2}{1-t}\right) \quad \text{for all } t \in ]0, 1[, \quad (2.2)$$

*and*

$$P\left[\frac{S_n}{n} > x\right] \leq \exp\left(-n\left(\sqrt{x+K} - \sqrt{K}\right)^2\right) \quad \text{for all } x > 0. \quad (2.3)$$

*Consequently,*

$$P\left[\frac{S_n}{n} > x\right] \leq \begin{cases} \exp\left(-\frac{nx^2}{K(1+\sqrt{2})^2}\right) & \text{if } x \in ]0, K[, \\ \exp\left(-\frac{nx}{(1+\sqrt{2})^2}\right) & \text{if } x \in ]K, \infty[. \end{cases} \quad (2.4)$$

*Conversely, if  $(X_k)$  are iid, and if  $P[\frac{S_n}{n} > x] \leq e^{-ncx}$  for some  $n > 1$ ,  $c > 0$ , and all  $x \geq x_1 > 0$  large enough, then for all  $\delta \in ]0, c[$ ,*

$$\mathbb{E}[e^{\delta X_1^+}] \leq K, \quad \text{where } X_1^+ = \max(X_1, 0), \text{ and } K = e^{\delta x_1} + \frac{\delta}{c-\delta} e^{-(c-\delta)x_1}.$$

**COROLLARY 2.2** *Under the conditions of Theorem 2.1,  $\forall \varepsilon > 0$ , there exist  $0 < x_0 < x_1$  and  $K_1 > 0$  depending only on  $K$  and  $\varepsilon$ , such that:*

$$P\left[\frac{S_n}{n} > x\right] \leq \begin{cases} \exp\left(-\frac{nx^2}{4K(1+\varepsilon)}\right) & \text{if } x \in ]0, x_0[, \\ \exp\left(-\frac{nx}{K_1}\right) & \text{if } x \in [x_0, x_1], \\ \exp\left(-\frac{nx}{1+\varepsilon}\right) & \text{if } x \in ]x_1, +\infty[. \end{cases} \quad (2.5)$$

We divide the proof into a series of lemmas.

**LEMMA 2.3** *Let  $(X_i)_{1 \leq i \leq n}$  be a finite sequence of random variables adapted to a filtration  $(\mathcal{F}_i)_{1 \leq i \leq n}$ . Let  $(l_i)_{1 \leq i \leq n}$  be a finite sequence of deterministic functions defined on a subinterval  $I$  of  $]0, \infty[$ , such that for each  $i$  and each  $t \in I$ ,*

$$\mathbb{E}[e^{tX_i} | \mathcal{F}_{i-1}] \leq e^{l_i(t)} \quad \text{a.s..} \quad (2.6)$$

*Then for every  $t \in I$ ,*

$$\mathbb{E}[e^{tS_n}] \leq \exp\left(\sum_{i=1}^n l_i(t)\right), \quad (2.7)$$

*and for every  $x > 0$ ,*

$$P\left[\frac{S_n}{n} > x\right] \leq e^{-nL_n^*(x)}, \quad (2.8)$$

*where*

$$L_n(t) = \frac{1}{n} \sum_{i=1}^n l_i(t), \quad \text{and } L_n^*(x) = \sup_{t \in I} (tx - L_n(t)). \quad (2.9)$$

**Proof.** (2.7) can be obtained by a simple induction argument on  $n$ . (2.8) is an immediate consequence of (2.7), since  $\forall x > 0, \forall t \in I$ ,

$$P\left[\frac{S_n}{n} > x\right] = P[e^{tS_n} > e^{tnx}] \leq e^{-ntx}\mathbb{E}[e^{tS_n}] \leq \exp(-n(tx - L_n(t))).$$

■

**REMARK 2.4** The submultiplicativity (2.7) for an adapted sequence corresponds to the multiplicativity  $\mathbb{E}[e^{tS_n}] = \prod_{i=1}^n \mathbb{E}[e^{tX_i}]$  in the independent case. This explains why it is natural to consider the conditional Laplace transform  $\mathbb{E}[e^{tX_i} | \mathcal{F}_{i-1}]$  in the supermartingale case, instead of the Laplace transform  $\mathbb{E}[e^{tX_i}]$  in the independent case. For example, using Lemma 2.3, we can obtain the following generalization of Petrov's inequality (p.54 of [31]):

**LEMMA 2.5** Let  $a_i > 0$  and  $T > 0$  be constants such that for all  $1 \leq i \leq n$  and all  $t \in ]0, T]$ , a.s.  $\mathbb{E}[e^{tX_i} | \mathcal{F}_{i-1}] \leq e^{a_i t^2}$ . Then for each  $A \geq \frac{1}{n} \sum_{i=1}^n a_i$ , we have

$$P\left[\frac{S_n}{n} > x\right] \leq \begin{cases} \exp\left(-\frac{nx^2}{4A}\right) & \text{if } x \in ]0, 2AT[, \\ \exp\left(-\frac{nTx}{2}\right) & \text{if } x \in [2AT, +\infty[. \end{cases} \quad (2.10)$$

**Proof.** We apply Lemma 2.3 with  $I = ]0, T]$  and  $l_i(t) = a_i t^2$ ,  $L_n(t) = At^2$ , which gives:

$$\mathbb{E}[e^{tS_n}] \leq \exp(nAt^2) \text{ for every } t \in ]0, T],$$

and

$$P\left[\frac{S_n}{n} > x\right] \leq e^{-nL_n^*(x)},$$

with  $L_n^*(x) = \sup_{t \in ]0, T]} (tx - At^2)$ . We calculate this sup and find:

$$L_n^*(x) = \begin{cases} \frac{x^2}{4A} & \text{if } x \in ]0, 2AT[, \\ Tx - AT^2 \geq \frac{Tx}{2} & \text{if } x \geq 2AT, \end{cases}$$

which ends the proof. ■

**LEMMA 2.6** Let  $X$  be a real-valued random variable defined on some probability space  $(\Omega, \mathcal{F}, \mathbb{P})$ , with  $\mathbb{E}X \leq 0$  and  $\mathbb{E}|e^{tX}| \leq K$  for some  $K > 0$ . Then for all  $t \in ]0, 1[$ ,

$$\mathbb{E}[e^{tX}] \leq \exp\left(\frac{Kt^2}{1-t}\right). \quad (2.11)$$

Consequently,

$$\mathbb{E}[e^{tX}] \leq \exp(2Kt^2) \text{ for every } t \in \left]0, \frac{1}{2}\right]. \quad (2.12)$$

**Proof.** Let  $t \in ]0, 1[$ . Since  $\mathbb{E}X \leq 0$ , we have

$$\begin{aligned} \mathbb{E}[e^{tX}] &= \sum_{k=0}^{\infty} t^k \mathbb{E}\left[\frac{X^k}{k!}\right] \leq 1 + \sum_{k=2}^{\infty} t^k \mathbb{E}\left[\frac{X^k}{k!}\right] \\ &\leq 1 + \sum_{k=2}^{\infty} t^k \mathbb{E}|e^{tX}| \leq 1 + K \frac{t^2}{1-t} \leq \exp\left(\frac{Kt^2}{1-t}\right). \end{aligned}$$

■

LEMMA 2.7 For  $K > 0$  and  $x > 0$ ,

$$\sup_{t \in ]0,1[} \left( tx - \frac{Kt^2}{1-t} \right) = \left( \sqrt{x+K} - \sqrt{K} \right)^2. \quad (2.13)$$

**Proof.** Let  $l_K(t) = \frac{Kt^2}{1-t}$ . We first consider  $l_1(t) = \frac{t^2}{1-t}$  (the case where  $K = 1$ ). Let  $h(t) = xt - \frac{t^2}{1-t}$ ,  $t \in ]0, 1[$ . Notice that  $h'(t) = 0$  if and only if  $x = \frac{t(2-t)}{(1-t)^2}$ , that is,  $t = 1 - \frac{1}{\sqrt{1+x}}$ . Therefore

$$l_1^*(x) = h \left( 1 - \frac{1}{\sqrt{1+x}} \right) = (\sqrt{x+1} - 1)^2.$$

In the general case, we have

$$l_K^*(x) = Kl_1^* \left( \frac{x}{K} \right) = \left( \sqrt{x+K} - \sqrt{K} \right)^2.$$

■

**Proof of Theorem 2.1.** By Lemma 2.6, we obtain that for every  $i$  and for every  $t \in ]0, 1[$ , a.s.

$$\mathbb{E}[e^{tX_i} | \mathcal{F}_{i-1}] \leq \exp \left( \frac{Kt^2}{1-t} \right).$$

Therefore by Lemmas 2.3 and 2.7, we obtain immediately (2.2) and (2.3).

To show (2.4), we notice that the function  $g(x) = \frac{(\sqrt{x+K}-\sqrt{K})^2}{x^2}$  is strictly decreasing on  $]0, +\infty[$  with  $\lim_{x \rightarrow +\infty} g(x) = 0$  and  $\lim_{x \rightarrow 0} g(x) = \frac{1}{4K}$ , whereas the function  $f(x) = \frac{(\sqrt{x+K}-\sqrt{K})^2}{x}$  is strictly increasing on  $]0, +\infty[$ , with  $\lim_{x \rightarrow +\infty} f(x) = 1$  and  $\lim_{x \rightarrow 0} f(x) = 0$ .

Therefore for every  $x \in ]0, K]$ ,  $(\sqrt{x+K} - \sqrt{K})^2 \geq x^2 g(K) = \frac{x^2}{K(1+\sqrt{2})^2}$ , and for every  $x > K$ ,  $(\sqrt{x+K} - \sqrt{K})^2 \geq xf(K) = \frac{x}{(1+\sqrt{2})^2}$ , which ends the proof of (2.4).

Conversely, suppose that  $(X_k)$  are iid, and that  $P \left[ \frac{S_n}{n} > x \right] \leq e^{-ncx}$  for some  $n > 1$ ,  $c > 0$  and all  $x \geq x_1 > 0$  large enough. Let  $\delta \in ]0, c[$ . Then for all  $x > 0$ ,

$$(P[X_1 > x])^n = P[X_i > x \text{ for all } 1 \leq i \leq n] \leq P \left[ \frac{S_n}{n} > x \right] \leq e^{-ncx},$$

so that  $P[X_1^+ > x] = P[X_1 > x] \leq e^{-cx}$ , and

$$\mathbb{E}[e^{\delta X_1^+}] = 1 + \int_0^{+\infty} P[X_1^+ > x] \delta e^{\delta x} dx \leq 1 + \int_0^{x_1} \delta e^{\delta x} dx + \int_{x_1}^{+\infty} \delta e^{-(c-\delta)x} dx = e^{\delta x_1} + \frac{\delta}{c-\delta} e^{-(c-\delta)x_1}.$$

■

**REMARK 2.8** Notice that by Lemma 2.6,  $\forall t \in ]0, \frac{1}{2}[$ , a.s.

$$\mathbb{E}[e^{tX_i} | \mathcal{F}_{i-1}] \leq \exp(2Kt^2).$$

Therefore by Lemma 2.5, we obtain immediately,

$$P \left[ \frac{S_n}{n} > x \right] \leq \begin{cases} \exp \left( -\frac{nx^2}{8K} \right) & \text{if } x \in ]0, 2K], \\ \exp \left( -\frac{nx}{4} \right) & \text{if } x > 2K. \end{cases} \quad (2.14)$$

But (2.3) of Theorem 2.1 gives more precise information.

**Proof of Corollary 2.2.** For  $\varepsilon \in ]0, 1[$ , let  $x_0 > 0$  and  $x_1 > 0$  be such that  $g(x_0) = \frac{1}{4K(1+\varepsilon)}$  and  $f(x_1) = \frac{1}{1+\varepsilon}$ , where  $g$  and  $f$  are as in the proof of Theorem 2.1.

If  $x \in ]0, x_0]$ , then  $(\sqrt{x+K} - \sqrt{K})^2 \geq x^2 g(x_0)$ , hence  $P[\frac{S_n}{n} > x] \leq \exp\left(-\frac{nx^2}{4K(1+\varepsilon)}\right)$ .

If  $x \in [x_1, +\infty[$ , then  $(\sqrt{x+K} - \sqrt{K})^2 \geq xf(x_1)$ , hence  $P[\frac{S_n}{n} > x] \leq \exp\left(-\frac{nx}{1+\varepsilon}\right)$ .

If  $x \in [x_0, x_1]$ , then  $(\sqrt{x+K} - \sqrt{K})^2 \geq xf(x_0) = xx_0g(x_0) = \frac{xx_0}{4K(1+\varepsilon)}$ . We set  $K_1 = \frac{4K(1+\varepsilon)}{x_0}$ , so that  $P[\frac{S_n}{n} > x] \leq \exp\left(-\frac{nx}{K_1}\right)$ .  $\blacksquare$

If we impose an exponential moment condition to  $X_i^2$  instead of  $X_i$ , we get the following Hoeffding type inequality.

**THEOREM 2.9** Let  $(X_i)_{1 \leq i \leq n}$  be a sequence of supermartingale differences adapted to  $(\mathcal{F}_i)$ . If there exist some constants  $R > 0$  and  $K > 0$  such that for all  $i$ ,

$$\mathbb{E}[e^{RX_i^2} | \mathcal{F}_{i-1}] \leq K \quad a.s., \quad (2.15)$$

then there exists a constant  $c > 0$  depending only on  $R$  and  $K$  such that:

$$\mathbb{E}[e^{tS_n}] \leq e^{ntc^2} \text{ for all } t > 0, \quad (2.16)$$

and

$$P\left[\frac{S_n}{n} > x\right] \leq e^{-\frac{nx^2}{4c}} \text{ for all } x > 0. \quad (2.17)$$

Conversely, if  $(X_i)$  are iid and if (2.16) or (2.17) holds for some  $n \geq 1$  and  $c > 0$ , then for each  $R \in ]0, \frac{1}{4c}[$ ,

$$\mathbb{E}[e^{RX_1^+}] \leq K, \quad \text{where } X_1^+ = \max(X_1, 0) \text{ and } K = 1 + \frac{R}{\frac{1}{4c} - R}.$$

Its proof will be based on the following Lemma.

**LEMMA 2.10** Let  $X$  be a random variable defined on a probability space  $(\Omega, \mathcal{F}, \mathbb{P})$ . If for some constants  $K$  and  $R > 0$ ,  $\mathbb{E}[e^{RX^2}] \leq K$ , then for all  $t > 0$ ,

$$\mathbb{E}[e^{t|X|}] \leq 1 + \frac{K\sqrt{\pi}}{2\sqrt{R}}t \exp\left(\frac{t^2}{4R}\right). \quad (2.18)$$

If additionally  $\mathbb{E}[X] \leq 0$ , then there exists  $a > 0$  depending only on  $K$  and  $R$  such that for all  $t > 0$ ,

$$\mathbb{E}[e^{tX}] \leq \exp\left(\frac{at^2}{2}\right). \quad (2.19)$$

**Proof.** By hypothesis  $P[|X| > x] \leq e^{-Rx^2} \mathbb{E}[e^{RX^2}] \leq Ke^{-Rx^2}$ . Hence for all  $t > 0$ ,

$$\begin{aligned} \mathbb{E}[e^{t|X|}] &= \int_0^{+\infty} P[e^{t|X|} > x] dx = \int_{-\infty}^{+\infty} P[|X| > u] d(e^{tu}) \\ &= 1 + t \int_0^{+\infty} P[|X| > u] e^{tu} du \leq 1 + Kt \int_0^{+\infty} e^{-Ru^2} e^{tu} du \\ &\leq 1 + \frac{K\sqrt{\pi}}{2\sqrt{R}}t \exp\left(\frac{t^2}{4R}\right). \end{aligned}$$

Let  $c > \frac{1}{4R}$ . Then there exists  $t_1 > 0$  such that

$$\forall t \geq t_1, \quad \mathbb{E}[e^{t|X|}] \leq \exp(ct^2). \quad (2.20)$$

On the other hand,

$$\mathbb{E}[e^{R|X|}] \leq \mathbb{E}[e^R; |X| \leq 1] + \mathbb{E}[e^{RX^2}; |X| > 1] \leq e^R + K,$$

so by Lemma 2.6, when  $\mathbb{E}[X] \leq 0$ , we have

$$\mathbb{E}[e^{tX}] \leq \exp\left(\frac{2K_1}{R^2}t^2\right) \quad \forall t \in \left[0, \frac{R}{2}\right], \quad (2.21)$$

where  $K_1 = e^R + K$ . From (2.20) and (2.21) we deduce that there exists  $a > 0$  depending only on  $K$  and  $R$ , such that

$$\forall t \geq 0, \quad \mathbb{E}[e^{tX}] \leq \exp\left(\frac{at^2}{2}\right). \quad \blacksquare$$

**Proof of Theorem 2.9.** Write  $\mathbb{E}_{i-1}[\cdot] = \mathbb{E}[\cdot | \mathcal{F}_{i-1}]$ . By Lemma 2.10 there exists  $a = a(R, K) > 0$  such that

$$\mathbb{E}_{i-1}[e^{tX_i}] \leq \exp\left(\frac{at^2}{2}\right) \quad \forall t > 0.$$

So by Lemmas 2.3 and 2.5, we get (2.16) and (2.17).

Conversely, suppose that  $(X_i)$  are iid and that (2.17) holds for some  $n \geq 1$  and  $c > 0$  (notice that (2.16) implies (2.17)). Let  $R \in ]0, \frac{1}{4c}[$ . Then  $\forall x > 0$ ,

$$(P[X_1 > x])^n = P[X_i > x \text{ for all } 1 \leq i \leq n] \leq P\left[\frac{S_n}{n} > x\right] \leq e^{-\frac{nx^2}{4c}},$$

so that  $P[X_1 > x] \leq e^{-\frac{x^2}{4c}}$ , and

$$\mathbb{E}[e^{RX_1^+}] = 1 + \int_0^{+\infty} P[X_1^+ > x] 2xRe^{Rx^2} dx = 1 + \int_0^{+\infty} P[X_1 > x] 2xRe^{Rx^2} dx \leq K,$$

where  $K = 1 + \int_0^{+\infty} 2xRe^{-(\frac{1}{4c}-R)x^2} dx = 1 + \frac{R}{\frac{1}{4c}-R}$ .  $\blacksquare$

### 3 Exponential bounds of $\mathbb{P}(S_n > nx)$ for large values of $x$

Notice that in the exponential inequality  $P(S_n \geq nx) \leq e^{-nc(x)}$  of the preceding section, for large  $x$ , we can take  $c(x) = cx$  or  $cx^2$  according to an exponential moment condition on  $X$  or on  $X^2$ , respectively. In this section we shall see that this property remains true for  $c(x) = cx^Q$  with any  $Q \geq 1$ .

**THEOREM 3.1** *Let  $(X_i)_{1 \leq i \leq n}$  be any adapted sequence with respect to a filtration  $(\mathcal{F}_i)_{1 \leq i \leq n}$ . Assume that there exist some constants  $Q > 1$ ,  $R > 0$  and  $K > 0$  such that for all  $i \in \llbracket 1, n \rrbracket$ ,*

$$\mathbb{E}[e^{R|X_i|^Q} | \mathcal{F}_{i-1}] \leq K \quad a.s.. \quad (3.1)$$

Let  $\rho > 1$  and  $\tau > 0$  be such that

$$\frac{1}{Q} + \frac{1}{\rho} = 1 \quad \text{and} \quad (\rho\tau)^{\frac{1}{\rho}}(QR)^{\frac{1}{Q}} = 1. \quad (3.2)$$

Then for any  $\tau_1 > \tau$ , there exists  $t_1 > 0$  depending only on  $K, Q, R$  and  $\tau_1$ , such that:

$$\mathbb{E}[e^{t|S_n|}] \leq \exp(n\tau_1 t^\rho) \quad \text{for all } t \geq t_1, \quad (3.3)$$

$$P\left[\frac{|S_n|}{n} > x\right] \leq \exp(-nR_1 x^Q) \quad \text{for all } x \geq x_1 := \rho\tau_1 t_1^{\rho-1}, \quad (3.4)$$

where  $R_1 > 0$  is such that  $(\rho\tau_1)^{\frac{1}{\rho}}(QR_1)^{\frac{1}{Q}} = 1$ .

Conversely, if  $(X_i)$  are iid and if (3.4) holds for some  $n \geq 1$ ,  $R_1 > 0$ ,  $Q > 1$  and  $x_1 > 0$ , then for all  $R \in ]0, R_1[$ ,

$$\mathbb{E}[e^{R|X_1|^Q}] \leq 2K, \quad \text{where } K = e^{Rx_1^Q} + \frac{R}{R_1 - R} e^{-(R_1 - R)x_1^Q}. \quad (3.5)$$

When  $(X_i)$  are supermartingale differences, we can complete Theorem 3.1 with an information for small values of  $x > 0$  and  $t > 0$ , as shown in the following theorem. The conclusion follows from Theorem 2.1 for small values of  $x, t > 0$ , and from Theorem 3.1 for large values of  $x, t > 0$ . The proof of (3.8) will be seen in the proof of (3.5). Notice that for large values of  $x, t > 0$ , the conclusion of Theorem 3.2 is sharper than that of Theorem 2.1.

**THEOREM 3.2** *Under the hypothesis of Theorem 3.1, if moreover  $(X_i)_{1 \leq i \leq n}$  is a sequence of supermartingale differences adapted to the filtration  $(\mathcal{F}_i)$ , then for any  $\tau_1 > \tau$ , there exist  $t_1 > 0$ ,  $x_1 > 0$ , and  $A, B > 0$ , depending only on  $K, Q, R$ , and  $\tau_1$ , such that:*

$$\mathbb{E}[e^{tS_n}] \leq \begin{cases} \exp(n\tau_1 t^\rho) & \text{if } t \geq t_1, \\ \exp(nAt^2) & \text{if } 0 \leq t \leq t_1, \end{cases} \quad (3.6)$$

and

$$P\left[\frac{S_n}{n} > x\right] \leq \begin{cases} \exp(-nR_1 x^Q) & \text{if } x \geq x_1, \\ \exp(-nBx^2) & \text{if } 0 \leq x \leq x_1. \end{cases} \quad (3.7)$$

Conversely, if  $(X_i)$  are iid and if the first inequality in (3.7) holds for some  $n \geq 1$ ,  $R_1 > 0$ ,  $Q > 1$  and  $x_1 > 0$ , then for all  $R \in ]0, R_1[$ ,

$$\mathbb{E}[e^{RX_1^+}] \leq K, \quad \text{where } X_1^+ = \max(X_1, 0) \text{ and } K = e^{Rx_1^Q} + \frac{R}{R_1 - R} e^{-(R_1 - R)x_1^Q}. \quad (3.8)$$

Before proving the theorems, we first give, for a positive random variable  $X$ , relations among the growth rate of the Laplace transform  $\mathbb{E}[e^{tX}]$  (as  $t \rightarrow \infty$ ), the decay rate of the tail probability  $P[X > x]$  (as  $x \rightarrow \infty$ ), and the exponential moments of the form  $\mathbb{E}[e^{RX^Q}]$  ( $Q > 1$ ).

**LEMMA 3.3** *(Relation between  $\mathbb{E}[e^{tX}]$  and  $P[X > x]$ )*

Let  $X$  be a positive real random variable. Let  $Q, \rho, \tau$ , and  $R \in ]0, +\infty[$  be such that  $1 < Q < +\infty$  and

$$\frac{1}{Q} + \frac{1}{\rho} = 1, \quad (\rho\tau)^{\frac{1}{\rho}}(QR)^{\frac{1}{Q}} = 1.$$

Let  $K > 0$  be a constant. Consider the following assertions:

- (1)  $\forall t > 0$ ,  $\mathbb{E}[e^{tX}] \leq Ke^{\tau t^\rho}$ ;
- (2)  $\forall x > 0$ ,  $P[X > x] \leq Ke^{-Rx^Q}$ ;

$$(3) \text{ For } a = K \left(\frac{2}{R}\right)^{\frac{1}{Q-1}} \text{ and all } t > 0, \mathbb{E}[e^{tX}] \leq 1 + K + at^\rho e^{\tau t^\rho}.$$

Then we have the following implications: (1)  $\Rightarrow$  (2)  $\Rightarrow$  (3).

Lemma 3.3 is closely related to the following Legendre duality between the functions  $t \mapsto \tau t^\rho$  and  $x \mapsto Rx^Q$ .

**LEMMA 3.4** *Let  $\rho > 1$ ,  $\tau > 0$  and  $t_0 \geq 0$ . Then  $\forall x \geq \rho\tau t_0^{\rho-1}$ ,*

$$\sup_{t \geq t_0} (tx - \tau t^\rho) = Rx^Q, \quad \text{where } \frac{1}{Q} + \frac{1}{\rho} = 1, (\rho\tau)^{\frac{1}{\rho}}(QR)^{\frac{1}{Q}} = 1.$$

**Proof.** The function  $h(t) = tx - \tau t^\rho$  attains its supremum on  $]0, +\infty[$  for  $t^* = (\frac{x}{\tau\rho})^{\frac{1}{\rho-1}}$ , and the supremum is  $h(t^*) = Rx^Q$ . As  $t^* \geq t_0$  if and only if  $x \geq \rho\tau t_0^{\rho-1}$ , we get the result.  $\blacksquare$

**Proof of Lemma 3.3.** We first prove the implication (1)  $\Rightarrow$  (2). If  $\mathbb{E}[e^{tX}] \leq Ke^{\tau t^\rho}$  then for every  $x > 0$  and  $t > 0$ ,

$$P[X > x] = P[e^{tX} > e^{tx}] \leq e^{-tx} \mathbb{E}[e^{tX}] \leq Ke^{\tau t^\rho - tx}.$$

Therefore by Lemma 3.4,  $P[X > x] \leq Ke^{-Rx^Q}$ .

We then prove the implication (2)  $\Rightarrow$  (3). If (2) holds, then for every  $t > 0$ ,

$$\mathbb{E}[e^{tX}] = 1 + t \int_0^{+\infty} P[X > x] e^{tx} dx \leq 1 + tK \int_0^{+\infty} e^{-Rx^Q + tx} dx.$$

We choose  $x_1 = (\frac{2t}{R})^{\frac{1}{Q-1}}$  so that  $-Rx^Q + tx \leq -xt$  for  $x \geq x_1$ ; by Lemma 3.4 (with  $t_0 = 0$ ),  $-Rx^Q + tx \leq \tau t^\rho$  for any  $x > 0$ . Therefore

$$\int_0^{+\infty} e^{-Rx^Q + tx} dx \leq \int_0^{x_1} e^{\tau t^\rho} dx + \int_{x_1}^{+\infty} e^{-xt} dx \leq x_1 e^{\tau t^\rho} + \frac{1}{t},$$

hence for  $a = K(\frac{2}{R})^{\frac{1}{Q-1}}$  and  $t > 0$ ,

$$\mathbb{E}[e^{tX}] \leq 1 + K + at^\rho e^{\tau t^\rho}.$$

■

**LEMMA 3.5** Let  $X$  be a positive real random variable. Let  $Q \in [1, +\infty[$ , and  $K, R \in ]0, +\infty[$ . Consider the following assertions:

- (1)  $\mathbb{E}[e^{RX^Q}] \leq K$ ;
- (2)  $\forall x > 0, P[X > x] \leq Ke^{-Rx^Q}$ ;
- (3) For any  $R_1 \in ]0, R[$ ,  $\mathbb{E}[e^{R_1 X^Q}] \leq \frac{R+R_1(K-1)}{R-R_1}$ .

Then we have the following implications: (1)  $\Rightarrow$  (2)  $\Rightarrow$  (3).

**Proof of Lemma 3.5.** The implication (1)  $\Rightarrow$  (2) is easy: if  $\mathbb{E}[e^{RX^Q}] \leq K$ , then  $P[X > x] = P[e^{RX^Q} > e^{Rx^Q}] \leq Ke^{-Rx^Q}$ . Let us now prove the implication (2)  $\Rightarrow$  (3). If  $P[X > x] \leq Ke^{-Rx^Q}$ , then for any  $R_1 \in ]0, R[$ ,

$$\begin{aligned} \mathbb{E}[e^{R_1 X^Q}] &= \int_0^{+\infty} P[e^{R_1 X^Q} > x] dx = 1 + R_1 Q \int_0^{+\infty} P[X > u] e^{R_1 u^Q} u^{Q-1} du \\ &\leq 1 + KR_1 Q \int_0^{+\infty} e^{(R_1 - R)u^Q} u^{Q-1} du = \frac{R+R_1(K-1)}{R-R_1}. \end{aligned}$$

■

**REMARK 3.6** Let  $Q, \rho \in ]0, +\infty[$  be such that  $1 < Q < +\infty$  and  $\frac{1}{Q} + \frac{1}{\rho} = 1$ . As a consequence of Lemma 3.3, we can easily see that writing

$$\begin{aligned} \tau &= \inf\{a > 0 : \mathbb{E}[e^{rX}] = O(\exp(ar^\rho))\}, \\ R &= \sup\{a > 0 : P[X > x] = O(\exp(-ax^Q))\}, \end{aligned}$$

we have

$$(\rho\tau)^{\frac{1}{\rho}}(QR)^{\frac{1}{Q}} = 1.$$

This was proved in a different way by Liu in [28]. It unifies Theorems 6.1, 7.1, 7.2, 7.3, 8.1, 9.1 and 9.2 of Ramachandran ([32]), and was first conjectured by Harris ([21]) in the context of branching processes.

**Proof of Theorem 3.1.** By Lemmas 3.5 and 3.3, we see that for  $a = K(\frac{2}{R})^{\frac{1}{Q-1}}$ ,

$$\mathbb{E}[e^{t|X_i|} | \mathcal{F}_{i-1}] \leq 1 + K + at^\rho e^{\tau t^\rho} \quad \forall t > 0.$$

Let  $\tau_1 > \tau$ . Then there exists  $t_1 > 0$  sufficiently large such that  $\forall t \geq t_1$ ,  $\mathbb{E}[e^{t|X_i|} | \mathcal{F}_{i-1}] \leq e^{\tau_1 t^\rho}$ . Applying Lemmas 2.3 and 3.4, we obtain that

$$\begin{aligned} \mathbb{E}[e^{t|S_n|}] &\leq \mathbb{E}[e^{t(|X_1| + \dots + |X_n|)}] \leq \exp(n\tau_1 t^\rho) \text{ if } t \geq t_1, \\ P\left[\frac{|S_n|}{n} > x\right] &\leq P\left[\frac{|X_1| + \dots + |X_n|}{n} > x\right] \leq e^{-nR_1 x^Q} \text{ if } x \geq x_1 = \rho\tau_1 t_1^{\rho-1}. \end{aligned}$$

Conversely, suppose that  $(X_k)$  are iid, and that  $P\left[\frac{|S_n|}{n} > x\right] \leq \exp(-nR_1 x^Q)$  for all  $x \geq x_1$ . Let  $R \in ]0, R_1[$ . Then for all  $x \geq x_1$ ,

$$(P[X_1 > x])^n = P[X_i > x \text{ for all } 1 \leq i \leq n] \leq P\left[\frac{S_n}{n} > x\right] \leq P\left[\frac{|S_n|}{n} > x\right] \leq \exp(-nR_1 x^Q),$$

so that  $X_1^+ = \max(0, X_1)$  satisfies  $P[X_1^+ > x] = P[X_1 > x] \leq e^{-R_1 x^Q}$ , and

$$\begin{aligned}\mathbb{E}[e^{R(X_1^+)^Q}] &= 1 + \int_0^{+\infty} P[X_1^+ > x] R Q x^{Q-1} e^{Rx^Q} dx \\ &\leq 1 + \int_0^{x_1} R Q x^{Q-1} e^{Rx^Q} dx + \int_{x_1}^{+\infty} R Q x^{Q-1} e^{-(R_1-R)x^Q} dx = e^{Rx_1^Q} + \frac{R}{R_1 - R} e^{-(R_1-R)x_1^Q}.\end{aligned}$$

By considering  $(-S_n)$  instead of  $(S_n)$ , we see that the same result holds for  $X_1^- = \max(0, -X_1)$ :

$$\mathbb{E}[e^{R(X_1^-)^Q}] \leq K := e^{Rx_1^Q} + \frac{R}{R_1 - R} e^{-(R_1-R)x_1^Q}.$$

Therefore

$$\mathbb{E}[e^{R|X_1|^Q}] \leq \mathbb{E}[e^{R(X_1^+)^Q}] + \mathbb{E}[e^{R(X_1^-)^Q}] \leq 2K.$$

■

**Proof of Theorem 3.2.** By Theorem 3.1, there exists  $t_1 > \frac{R}{2}$  such that

$$\mathbb{E}[e^{tS_n}] \leq \mathbb{E}[e^{t|S_n|}] \leq \exp(n\tau_1 t^\rho) \text{ for all } t \geq t_1,$$

and

$$P\left[\frac{S_n}{n} > x\right] \leq P\left[\frac{|S_n|}{n} > x\right] \leq \exp(-nR_1 x^Q) \text{ for all } x \geq x_1 := \rho\tau_1 t_1^{\rho-1}.$$

On the other hand, notice that  $\mathbb{E}[e^{R|X_i|} | \mathcal{F}_{i-1}] \leq K_1 := e^R + K$ , so that by Theorem 2.1,

$$\mathbb{E}[e^{tS_n}] \leq \exp\left(\frac{2nK_1 t^2}{R^2}\right) \text{ for all } t \in \left[0, \frac{R}{2}\right].$$

If  $t \in \left]\frac{R}{2}, t_1\right]$ , then

$$\mathbb{E}[e^{tS_n}] \leq \mathbb{E}[e^{t_1|S_n|}] \leq \exp(n\tau_1 t_1^\rho) \leq e^{n\frac{4\tau_1 t_1^\rho}{R^2} t^2}.$$

Set  $A = \max(\frac{2K_1}{R^2}, \frac{4\tau_1 t_1^\rho}{R^2})$ . Then

$$\mathbb{E}[e^{tS_n}] \leq e^{nAt^2} \quad \forall t \in [0, t_1].$$

Again by Theorem 2.1, we can choose  $B > 0$  small enough such that

$$P\left[\frac{S_n}{n} > x\right] \leq \exp(-nBx^2) \text{ if } x \in [0, x_1].$$

■

## 4 Extension to the case $\mathbb{E}[e^{|X_i|} | \mathcal{F}_{i-1}] \leq K_i$

The following theorems are immediate generalizations of Theorems 2.1, 2.9, 3.1 and 3.2. The proofs of the first two theorems remain the same; the proof of the third needs a short argument for the concerned constants to be independent of  $n$ . The first theorem is an extension of Bernstein's inequality.

**THEOREM 4.1** *Let  $(X_i)_{1 \leq i \leq n}$  be a finite sequence of supermartingale differences. If for some constants  $K_i > 0$  and all  $i \in \llbracket 1, n \rrbracket$ , a.s.*

$$\mathbb{E}[e^{|X_i|} | \mathcal{F}_{i-1}] \leq K_i, \tag{4.1}$$

*then for each  $K \geq \frac{K_1 + \dots + K_n}{n}$ ,*

$$\mathbb{E}[e^{tS_n}] \leq \exp\left(\frac{nKt^2}{1-t}\right) \text{ for all } t \in [0, 1[, \tag{4.2}$$

and

$$P\left[\frac{S_n}{n} > x\right] \leq \exp\left(-n\left(\sqrt{x+K} - \sqrt{K}\right)^2\right) \text{ for all } x > 0. \tag{4.3}$$

Consequently,

$$P\left[\frac{S_n}{n} > x\right] \leq \begin{cases} \exp\left(-\frac{nx^2}{K(1+\sqrt{2})^2}\right) & \text{if } x \in [0, K], \\ \exp\left(-\frac{nx}{(1+\sqrt{2})^2}\right) & \text{if } x > K. \end{cases} \quad (4.4)$$

The second theorem is an extension of Hoeffding's inequality.

**THEOREM 4.2** Let  $(X_i)_{1 \leq i \leq n}$  be a sequence of supermartingale differences adapted to  $(\mathcal{F}_i)$ . If there exist some constants  $K_i > 0$  and  $R > 0$  such that for all  $i \in [1, n]$ , a.s.

$$\mathbb{E}[e^{RX_i^2} | \mathcal{F}_{i-1}] \leq K_i, \quad (4.5)$$

then for each  $K \geq \frac{K_1 + \dots + K_n}{n}$ , there exists a constant  $c > 0$  depending only on  $R$  and  $K$  such that:

$$\mathbb{E}[e^{tS_n}] \leq e^{nt^2} \text{ for all } t > 0, \quad (4.6)$$

and

$$P\left[\frac{S_n}{n} > x\right] \leq e^{-\frac{nx^2}{4c}} \text{ for all } x > 0. \quad (4.7)$$

The third theorem shows a close relation between  $P[|X_i| > x]$  and  $P\left[\frac{|S_n|}{n} > x\right]$  for large values of  $x > 0$ . Notice that this result is valid for any adapted sequence.

**THEOREM 4.3** Let  $(X_i)_{1 \leq i \leq n}$  be any adapted sequence with respect to a filtration  $(\mathcal{F}_i)_{1 \leq i \leq n}$ . Assume that there exist some constants  $Q > 1$ ,  $R > 0$  and  $K_i > 0$  such that for all  $i \in [1, n]$ ,

$$\mathbb{E}[e^{R|X_i|^Q} | \mathcal{F}_{i-1}] \leq K_i \quad \text{a.s..} \quad (4.8)$$

Let  $\rho > 1$  and  $\tau > 0$  be such that

$$\frac{1}{Q} + \frac{1}{\rho} = 1 \text{ and } (\rho\tau)^{\frac{1}{\rho}}(QR)^{\frac{1}{Q}} = 1. \quad (4.9)$$

Let  $K \geq \frac{K_1 + \dots + K_n}{n}$ . Then for any  $\tau_1 > \tau$ , there exists  $t_1 > 0$  depending only on  $K, Q, R$  and  $\tau_1$ , such that:

$$\mathbb{E}[e^{t|S_n|}] \leq \exp(n\tau_1 t^\rho) \text{ for all } t \geq t_1, \quad (4.10)$$

and

$$P\left[\frac{|S_n|}{n} > x\right] \leq \exp(-nR_1 x^Q) \text{ for all } x \geq x_1 := \rho\tau_1 t_1^{\rho-1}, \quad (4.11)$$

where  $R_1$  is such that  $(\rho\tau_1)^{\frac{1}{\rho}}(QR_1)^{\frac{1}{Q}} = 1$ .

**Proof.** By Lemmas 3.5 and 3.3, we see that for  $a = (\frac{2}{R})^{\frac{1}{Q-1}}$ ,

$$\mathbb{E}[e^{t|X_i|} | \mathcal{F}_{i-1}] \leq 1 + K_i(1 + at^\rho e^{\tau t^\rho}) \quad \forall t > 0.$$

By Lemma 2.3,

$$\mathbb{E}[e^{t|S_n|}] \leq \mathbb{E}[e^{t(|X_1| + \dots + |X_n|)}] \leq \prod_{i=1}^n (1 + K_i(1 + at^\rho e^{\tau t^\rho})).$$

It is easy to see that  $1 + K_i(1 + at^\rho e^{\tau t^\rho}) \leq e^{K_i}(1 + at^\rho e^{\tau t^\rho})$ , so we have

$$\mathbb{E}[e^{t|S_n|}] \leq \left(e^K(1 + at^\rho e^{\tau t^\rho})\right)^n.$$

Let  $\tau_1 > \tau$ . Then there exists  $t_1 > 0$  sufficiently large such that  $\forall t \geq t_1$ ,  $e^K(1 + at^\rho e^{\tau t^\rho}) \leq e^{\tau_1 t^\rho}$ , which gives (4.10). As

$$P\left[\frac{|S_n|}{n} > x\right] = P[e^{t|S_n|} > e^{tnx}] \leq e^{-ntx} \mathbb{E}[e^{t|S_n|}] \leq \exp(-n(tx - \tau_1 t^\rho)),$$

we deduce (4.11) from Lemma 3.4. ■

As in section 3, when  $(X_i)$  are supermartingale differences, using Theorem 4.1 we can complete Theorem 4.3 with an information for small values of  $x > 0$  and  $t > 0$ , as shown in the following theorem. For large values of  $x, t > 0$ , it gives inequalities sharper than those of Theorem 4.1.

**THEOREM 4.4** *Under the same hypothesis as in Theorem 4.3, if moreover  $(X_i)_{1 \leq i \leq n}$  is a sequence of supermartingale differences adapted to the filtration  $(\mathcal{F}_i)$ , then for any  $\tau_1 > \tau$ , there exist  $t_1 > 0$ ,  $x_1 > 0$ , and  $A, B > 0$ , depending only on  $K, Q, R$  and  $\tau_1$ , such that:*

$$\mathbb{E}[e^{tS_n}] \leq \begin{cases} \exp(n\tau_1 t^Q) & \text{if } t \geq t_1, \\ \exp(nAt^2) & \text{if } 0 \leq t \leq t_1, \end{cases} \quad (4.12)$$

and

$$P\left[\frac{S_n}{n} > x\right] \leq \begin{cases} \exp(-nR_1 x^Q) & \text{if } x \geq x_1, \\ \exp(-nBx^2) & \text{if } 0 \leq x \leq x_1. \end{cases} \quad (4.13)$$

## 5 Rate of convergence with probability 1 and in $L^p$

**THEOREM 5.1** *Let  $(X_i)_{1 \leq i \leq n}$  be a sequence of supermartingale differences. If for some constants  $K_i > 0$  and for all  $i \in \llbracket 1, n \rrbracket$ ,*

$$\mathbb{E}[e^{|X_i|} | \mathcal{F}_{i-1}] \leq K_i \quad \text{a.s.}, \quad (5.1)$$

*then writing  $K = \limsup_{n \rightarrow +\infty} \frac{K_1 + \dots + K_n}{n}$  and  $S_n^+ = \max(0, S_n)$ , we have:*

$$\limsup_{n \rightarrow +\infty} \frac{S_n^+}{\sqrt{n \ln n}} \leq 2\sqrt{K} \quad \text{a.s.}, \quad (5.2)$$

and for every  $p > 0$ ,

$$\limsup_{n \rightarrow +\infty} n^{\frac{p}{2}} \mathbb{E}\left[\left(\frac{S_n^+}{n}\right)^p\right] \leq p 2^{p-1} K^{\frac{p}{2}} \Gamma\left(\frac{p}{2}\right). \quad (5.3)$$

**Proof.** For the proof of (5.2), by Borel-Cantelli's Lemma, it suffices to show that for every  $a > 2\sqrt{K}$ ,

$$\sum_{n=0}^{+\infty} P\left[\frac{S_n^+}{\sqrt{n \ln n}} > a\right] < +\infty.$$

Let us fix  $a > 2\sqrt{K}$ . Let  $\varepsilon > 0$  be such that  $a > 2\sqrt{K + \varepsilon}$  and let  $n_1 > 0$  be such that for every  $n \geq n_1$ ,  $\frac{K_1 + \dots + K_n}{n} < K + \varepsilon$ . Then we deduce from Theorem 4.1 that for every  $n \geq n_1$ ,

$$P\left[\frac{S_n^+}{\sqrt{n \ln n}} > a\right] = P\left[\frac{S_n}{n} > \sqrt{\frac{\ln n}{n}} a\right] \leq \exp\left(-n\left(\sqrt{x_n + K + \varepsilon} - \sqrt{K + \varepsilon}\right)^2\right),$$

with  $x_n = \sqrt{\frac{\ln n}{n}} a$ . When  $n$  tends to  $\infty$ ,  $n\left(\sqrt{x_n + K + \varepsilon} - \sqrt{K + \varepsilon}\right)^2 \sim \frac{a^2 \ln n}{4(K + \varepsilon)}$ . As  $a^2 > 4(K + \varepsilon)$ , it follows that

$$\sum_{n=0}^{+\infty} P\left[\frac{S_n^+}{\sqrt{n \ln n}} > a\right] < +\infty.$$

This ends the proof of (5.2).

We now come to the proof of (5.3). Let  $n_1 > 0$  be as in the proof of (5.2). We deduce from Theorem 4.1 that for every  $n \geq n_1$ ,

$$\mathbb{E}\left[\left(\frac{S_n^+}{n}\right)^p\right] = p \int_0^\infty P\left[\frac{S_n}{n} > x\right] x^{p-1} dx \leq p \int_0^\infty \exp\left(-n\left(\sqrt{x + K + \varepsilon} - \sqrt{K + \varepsilon}\right)^2\right) x^{p-1} dx.$$

Set  $y = n\left(\sqrt{x + K + \varepsilon} - \sqrt{K + \varepsilon}\right)^2$ . Then  $\sqrt{\frac{y}{n}} = \sqrt{x + K + \varepsilon} - \sqrt{K + \varepsilon}$ ,  $x = \sqrt{\frac{y}{n}}(\sqrt{\frac{y}{n}} + 2\sqrt{K + \varepsilon})$ ,  $dx = \frac{\sqrt{\frac{y}{n}} + \sqrt{K + \varepsilon}}{\sqrt{n}\sqrt{y}} dy$ , so that

$$\mathbb{E}\left[\left(\frac{S_n^+}{n}\right)^p\right] \leq \frac{c_n(p)}{n^{\frac{p}{2}}}, \quad (5.4)$$

where

$$c_n(p) = p \int_0^\infty e^{-y} y^{\frac{p}{2}-1} \left( \sqrt{\frac{y}{n}} + 2\sqrt{K+\varepsilon} \right)^{p-1} \left( \sqrt{\frac{y}{n}} + \sqrt{K+\varepsilon} \right) dy$$

satisfies

$$\lim_{n \rightarrow \infty} c_n(p) = p \int_0^\infty e^{-y} y^{\frac{p}{2}-1} \left( 2\sqrt{K+\varepsilon} \right)^{p-1} \left( \sqrt{K+\varepsilon} \right) dy = p 2^{p-1} (K+\varepsilon)^{\frac{p}{2}} \Gamma\left(\frac{p}{2}\right).$$

■

In the case of a sequence of martingale differences, replacing  $S_n^+$  by  $|S_n|$  in the proof above, we obtain immediately:

**COROLLARY 5.2** *Let  $(X_i)_{1 \leq i \leq n}$  be a sequence of martingale differences. If for some constants  $K_i > 0$  and for all  $i \in \llbracket 1, n \rrbracket$ ,*

$$\mathbb{E}[e^{|X_i|} | \mathcal{F}_{i-1}] \leq K_i \quad a.s., \quad (5.5)$$

*then for  $K = \limsup_{n \rightarrow +\infty} \frac{K_1 + \dots + K_n}{n}$ ,*

$$\limsup_{n \rightarrow +\infty} \frac{|S_n|}{\sqrt{n \ln n}} \leq 2\sqrt{K} \quad a.s., \quad (5.6)$$

*and for every  $p > 0$ ,*

$$\limsup_{n \rightarrow +\infty} n^{\frac{p}{2}} \mathbb{E} \left[ \left( \frac{|S_n|}{n} \right)^p \right] \leq p 2^p K^{\frac{p}{2}} \Gamma\left(\frac{p}{2}\right). \quad (5.7)$$

**REMARK 5.3** The exponential moment condition (5.5) can certainly be relaxed for a result of the form  $\mathbb{E} \left[ \left( \frac{|S_n|}{n} \right)^p \right] = O(n^{-\frac{p}{2}})$ . For example, as shown in [27], p.150, by Burkholder's inequality, we can obtain the following result: if  $p \geq 2$  and  $\mathbb{E}[|X_i|^p] \leq K$  for some  $K > 0$  and all  $i \in \llbracket 1, n \rrbracket$ , then

$$\mathbb{E}[|S_n|^p] \leq n^{\frac{p}{2}} (18pq^{1/2})^p K, \quad (5.8)$$

where  $\frac{1}{p} + \frac{1}{q} = 1$ .

## 6 Free energy of directed polymers: concentration inequalities

We now consider the model of a directed polymer in a random environment, already described in the introduction. For convenience, let us recall it briefly as follows. Let  $\omega = (\omega_n)_{n \in \mathbb{N}}$  be the simple random walk on the  $d$ -dimensional integer lattice  $\mathbb{Z}^d$  starting at 0, defined on a probability space  $(\Omega, \mathcal{F}, P)$ . Let  $\eta = (\eta(n, x))_{(n, x) \in \mathbb{N} \times \mathbb{Z}^d}$  be a sequence of real valued, non constant and i.i.d. random variables defined on another probability space  $(E, \mathcal{E}, \mathbb{Q})$ . The path  $\omega$  represents the directed polymer and  $\eta$  the random environment. For any  $n > 0$ , define the random polymer measure  $\mu_n$  on the path space  $(\Omega, \mathcal{F})$  by

$$\mu_n = \frac{1}{Z_n(\beta)} \exp(\beta H_n(\omega)) P(d\omega), \quad (6.1)$$

where  $\beta \in \mathbb{R}$  is the inverse temperature,

$$H_n(\omega) = \sum_{j=1}^n \eta(j, \omega_j), \text{ and } Z_n(\beta) = P[\exp(\beta H_n(\omega))]. \quad (6.2)$$

Let  $\lambda(\beta) = \ln \mathbb{Q}[e^{\beta \eta(0,0)}]$  be the logarithmic moment generating function of  $\eta(0,0)$ . We fix  $\beta > 0$  (otherwise we consider  $-\eta$ ), and assume only  $\lambda(\pm\beta) < \infty$ , which is equivalent to  $\mathbb{Q}[e^{\beta |\eta(0,0)|}] < \infty$ . We are interested in the asymptotic behaviour of the normalized partition function

$$W_n(\beta) = \frac{Z_n(\beta)}{\mathbb{Q}[Z_n(\beta)]} = P[\exp(\beta H_n - n\lambda(\beta))], \quad (6.3)$$

and the free energy  $\frac{1}{n} \ln W_n(\beta)$ . For simplicity, we shall write  $W_n$  for  $W_n(\beta)$ ,  $Z_n$  for  $Z_n(\beta)$ , and  $\eta$  for  $\eta(0,0)$ . We use the same letter  $\eta$  to denote the environment sequence  $(\eta(n,x))_{(n,x) \in \mathbb{N} \times \mathbb{Z}^d}$  and the random variable  $\eta(0,0)$ ; there will be no confusion according to the context. In this section, we shall prove exponential concentration inequalities for the free energies  $\frac{\ln W_n}{n}$ , and convergence results of the centered energies  $\frac{\ln W_n}{n} - \frac{\mathbb{Q}[\ln W_n]}{n}$ : cf. Theorems 6.1, 6.5, 6.6, and their corollaries.

**THEOREM 6.1** *Assume that  $\mathbb{Q}[e^{\beta|\eta|}] < +\infty$ , and set  $K = 2 \exp(\lambda(-\beta)) + \lambda(\beta)$ . Then for all  $n \geq 1$ ,*

$$\mathbb{Q}[e^{\pm t(\ln W_n - \mathbb{Q}[\ln W_n])}] \leq \exp\left(\frac{nKt^2}{1-t}\right) \text{ for all } t \in ]0, 1[, \quad (6.4)$$

and

$$\mathbb{Q}\left[\pm \frac{1}{n}(\ln W_n - \mathbb{Q}[\ln W_n]) > x\right] \leq \exp\left(-n\left(\sqrt{x+K} - \sqrt{K}\right)^2\right) \text{ for all } x > 0. \quad (6.5)$$

Consequently,  $\forall n \geq 1$ ,

$$\mathbb{Q}\left[\pm \frac{1}{n}(\ln W_n - \mathbb{Q}[\ln W_n]) > x\right] \leq \begin{cases} \exp\left(-\frac{nx^2}{K(1+\sqrt{2})^2}\right) & \text{if } x \in ]0, K], \\ \exp\left(-\frac{nx}{(1+\sqrt{2})^2}\right) & \text{if } x > K. \end{cases} \quad (6.6)$$

**COROLLARY 6.2** *Under the conditions of Theorem 6.1,  $\forall \varepsilon > 0$ , there exist  $0 < x_0 < x_1$  and  $K_1 > 0$  depending only on  $K$  and  $\varepsilon$ , such that:*

$$\mathbb{Q}\left[\pm \frac{1}{n}(\ln W_n - \mathbb{Q}[\ln W_n]) > x\right] \leq \begin{cases} \exp\left(-\frac{nx^2}{4K(1+\varepsilon)}\right) & \text{if } x \in ]0, x_0[, \\ \exp\left(-\frac{nx}{K_1}\right) & \text{if } x \in [x_0, x_1], \\ \exp\left(-\frac{nx}{1+\varepsilon}\right) & \text{if } x \in ]x_1, +\infty[. \end{cases} \quad (6.7)$$

**REMARK 6.3** Using Lesigne and Volny's martingale inequality (1.5), Comets, Shiga and Yoshida (2003, [11]) proved that if  $\mathbb{Q}[e^{\beta|\eta|}] < +\infty$  for all  $\beta > 0$ , then  $\forall x > 0$ , there exists  $n_0 \in \mathbb{N}^*$  such that for any  $n \geq n_0$ ,

$$\mathbb{Q}\left[\left|\frac{1}{n} \ln W_n - \frac{1}{n} \mathbb{Q}[\ln W_n]\right| > x\right] \leq \exp\left(-\frac{n^{\frac{1}{3}}x^{\frac{2}{3}}}{4}\right). \quad (6.8)$$

Our result is sharper as  $n^{1/3}$  is replaced by  $n$ . Another advantage is that our conclusion holds for all  $n$ , not only for  $n$  large enough; thanks to this advantage, we can use our inequalities to study the convergence rate for the a.s. and  $L^p$  convergence: cf. Theorem 6.5. The third advantage is that we assume  $\mathbb{Q}[e^{\beta|\eta|}] < +\infty$  only for the fixed  $\beta$ , not for all  $\beta > 0$ . The first two advantages are due to the application of our exponential martingale inequality (Theorem 2.1); the third one comes from a direct estimation of the conditional exponential moment (Lemma 6.4) by use of convex inequalities, without using Lemma 3.1 of [11].

For the proof, as in [11], we write  $\ln W_n - \mathbb{Q}[\ln W_n]$  as a sum of  $(\mathcal{E}_j)_{1 \leq j \leq n}$  martingale differences:

$$\ln W_n - \mathbb{Q}[\ln W_n] = \sum_{j=1}^n V_{n,j}, \text{ with } V_{n,j} = \mathbb{Q}_j[\ln W_n] - \mathbb{Q}_{j-1}[\ln W_n],$$

where  $\mathbb{Q}_j$  denotes the conditional expectation with respect to  $\mathbb{Q}$  given  $\mathcal{E}_j$ ,  $\mathcal{E}_j = \sigma[\eta(i, x) : 1 \leq i \leq j, x \in \mathbb{Z}^d]$ .

**LEMMA 6.4** *We have*

$$\mathbb{Q}_{j-1}[\exp(tV_{n,j})] \leq \exp(L(t)) \text{ for every } t \in \mathbb{R}, \quad (6.9)$$

where

$$L(t) = \begin{cases} \lambda(t\beta) + \lambda(-t\beta) & \text{if } |t| > 1, \\ \lambda(-|t|\beta) + |t|\lambda(\beta) & \text{if } |t| \leq 1. \end{cases} \quad (6.10)$$

Consequently,

$$\mathbb{Q}_{j-1}[\exp(|V_{n,j}|)] \leq K := 2 \exp(\lambda(\beta) + \lambda(-\beta)). \quad (6.11)$$

**Proof.** We fix  $t \in \mathbb{R}^*$  and assume  $L(t) < \infty$  (otherwise there is nothing to prove). Set

$$e_{n,j} = \exp \left( \sum_{1 \leq k \leq n, k \neq j} (\beta \eta(k, \omega_k) - \lambda(\beta)) \right), \quad W_{n,j} = P[e_{n,j}].$$

Since  $\mathbb{Q}_{j-1}[\ln W_{n,j}] = \mathbb{Q}_j[\ln W_{n,j}]$ , we have

$$V_{n,j} = \mathbb{Q}_j \left[ \ln \frac{W_n}{W_{n,j}} \right] - \mathbb{Q}_{j-1} \left[ \ln \frac{W_n}{W_{n,j}} \right]. \quad (6.12)$$

For  $j \in \mathbb{N}$  and  $x \in \mathbb{Z}^d$ , define

$$\bar{\eta}_x = \bar{\eta}(j, x) = \exp(\beta \eta(j, x) - \lambda(\beta)), \quad \alpha_x = \frac{P[e_{n,j}; \omega_j = x]}{W_{n,j}}.$$

(Throughout the paper, for a measure  $\mu$ , a function  $f$ , and a set  $A$ , we use the notation  $\mu[f; A] = \int f \mathbf{1}_A d\mu$ , where  $\mathbf{1}_A$  is the indicator function of  $A$ ). Then

$$\sum_{x \in \mathbb{Z}^d} \alpha_x = 1 \text{ and } \frac{W_n}{W_{n,j}} = \sum_{x \in \mathbb{Z}^d} \alpha_x \bar{\eta}_x.$$

By (6.12),

$$\mathbb{Q}_{j-1}[\exp(tV_{n,j})] = \exp \left( -t \mathbb{Q}_{j-1} \left[ \ln \frac{W_n}{W_{n,j}} \right] \right) \mathbb{Q}_{j-1} \left[ \exp \left( t \mathbb{Q}_j \left[ \ln \frac{W_n}{W_{n,j}} \right] \right) \right].$$

Since the function  $x \mapsto e^{tx}$  is convex, using Jensen's inequality and the fact that  $\mathcal{E}_{j-1} \subset \mathcal{E}_j$ , we get:

$$\mathbb{Q}_{j-1}[\exp(tV_{n,j})] \leq \mathbb{Q}_{j-1} \left[ \left( \frac{W_n}{W_{n,j}} \right)^{-t} \right] \mathbb{Q}_{j-1} \left[ \left( \frac{W_n}{W_{n,j}} \right)^t \right]. \quad (6.13)$$

If  $t < 0$  or  $t \geq 1$  then the function  $x \mapsto x^t$  is convex, therefore by Jensen's inequality we have

$$\left( \frac{W_n}{W_{n,j}} \right)^t = \left( \sum_{x \in \mathbb{Z}^d} \alpha_x \bar{\eta}_x \right)^t \leq \sum_{x \in \mathbb{Z}^d} \alpha_x (\bar{\eta}_x)^t.$$

We consider the  $\sigma$ -algebra  $\mathcal{E}_{n,j} = \sigma[\eta(k, x); 1 \leq k \leq n, k \neq j, x \in \mathbb{Z}^d]$ . Then  $\mathcal{E}_{j-1} \subset \mathcal{E}_{n,j}$ , the  $\alpha_x$  are  $\mathcal{E}_{n,j}$ -measurable, and the  $\bar{\eta}_x$  are independent of  $\mathcal{E}_{n,j}$ , so that

$$\mathbb{Q}_{j-1}[\alpha_x (\bar{\eta}_x)^t] = \mathbb{Q}_{j-1}[\mathbb{Q}[\alpha_x (\bar{\eta}_x)^t | \mathcal{E}_{n,j}]] = \mathbb{Q}_{j-1}[\alpha_x \mathbb{Q}[(\bar{\eta}_x)^t]] = \exp(\lambda(t\beta) - t\lambda(\beta)) \mathbb{Q}_{j-1}[\alpha_x].$$

Hence for  $t < 0$  or  $t \geq 1$ ,

$$\mathbb{Q}_{j-1} \left[ \left( \frac{W_n}{W_{n,j}} \right)^t \right] \leq \exp(\lambda(t\beta) - t\lambda(\beta)). \quad (6.14)$$

It is easily seen that the equality holds for  $t = 1$ :  $\mathbb{Q}_{j-1} \left[ \frac{W_n}{W_{n,j}} \right] = 1$ . Again by Jensen's inequality, we have, for  $t \in ]0, 1]$ ,

$$\mathbb{Q}_{j-1} \left[ \left( \frac{W_n}{W_{n,j}} \right)^t \right] \leq \left( \mathbb{Q}_{j-1} \left[ \frac{W_n}{W_{n,j}} \right] \right)^t = 1. \quad (6.15)$$

The inequality (6.9) is then just a combination of (6.13), (6.14), and (6.15). In particular,

$$\mathbb{Q}_{j-1}[\exp(\pm V_{n,j})] \leq \exp(\lambda(\beta) + \lambda(-\beta)), \text{ so that } \mathbb{Q}_{j-1}[\exp(|V_{n,j}|)] \leq K := 2 \exp(\lambda(\beta) + \lambda(-\beta)).$$

■

**Proof of Theorem 6.1.** From Lemma 6.4 and Theorem 2.1, we deduce:

$$\mathbb{Q}[e^{t(\ln W_n - \mathbb{Q}[\ln W_n])}] \leq \exp\left(\frac{nKt^2}{1-t}\right) \text{ for all } t \in ]0, 1[, \quad (6.16)$$

and

$$\mathbb{Q}\left[\frac{1}{n}(\ln W_n - \mathbb{Q}[\ln W_n]) > x\right] \leq \exp\left(-n\left(\sqrt{x+K} - \sqrt{K}\right)^2\right) \text{ for all } x > 0. \quad (6.17)$$

Applying Theorem 2.1 to the sequence  $(-V_{n,j})$ , we find that

$$\mathbb{Q}[e^{-t(\ln W_n - \mathbb{Q}[\ln W_n])}] \leq \exp\left(\frac{nKt^2}{1-t}\right) \text{ for every } t \in ]0, 1[, \quad (6.18)$$

and

$$\mathbb{Q}\left[-\frac{1}{n}(\ln W_n - \mathbb{Q}[\ln W_n]) > x\right] \leq \exp\left(-n\left(\sqrt{x+K} - \sqrt{K}\right)^2\right) \text{ for all } x > 0. \quad (6.19)$$

The inequalities (6.16) and (6.18) give (6.4), (6.17) and (6.19) give (6.5).  $\blacksquare$

**Proof of Corollary 6.2.** The proof is the same as the proof of Corollary 2.2.  $\blacksquare$

**THEOREM 6.5** Assume that  $\mathbb{Q}[e^{\beta|\eta|}] < +\infty$ , and set  $K = 2 \exp(\lambda(-\beta)) + \lambda(\beta)$ . Then

$$\frac{1}{n} \ln W_n - \frac{1}{n} \mathbb{Q}[\ln W_n] \rightarrow 0 \text{ a.s. and in } L^p, \quad (6.20)$$

with

$$\limsup_{n \rightarrow +\infty} \sqrt{\frac{n}{\ln n}} \left| \frac{\ln W_n}{n} - \frac{\mathbb{Q}[\ln W_n]}{n} \right| \leq 2\sqrt{K} \text{ a.s.}, \quad (6.21)$$

and for every  $p > 0$ ,

$$\limsup_{n \rightarrow +\infty} n^{\frac{p}{2}} \mathbb{Q}\left[\left| \frac{\ln W_n - \mathbb{Q}[\ln W_n]}{n} \right|^p\right] \leq p 2^p K^{\frac{p}{2}} \Gamma\left(\frac{p}{2}\right). \quad (6.22)$$

**Proof.** Recall that with the notations of the proof of Theorem 6.1, we have

$$\mathbb{Q}_{j-1}[\exp(|V_{n,j}|)] \leq K.$$

Then the inequalities (6.21) and (6.22) are consequences of the inequalities (5.6) and (5.7) of Corollary 5.2.  $\blacksquare$

**THEOREM 6.6** Assume that  $K_0 := \mathbb{Q}[e^{R|\eta|^Q}] < +\infty$  for some  $Q > 1$  and  $R > 0$ . Let  $\rho > 1$  and  $\tau > 0$  be determined by

$$\frac{1}{Q} + \frac{1}{\rho} = 1 \text{ and } (\rho\tau)^{\frac{1}{\rho}}(QR)^{\frac{1}{Q}} = 1. \quad (6.23)$$

Then for each  $\tau_1 > \tau$ , there exist constants  $t_0, A, B > 0$ , depending only on  $\beta, K_0, Q, R, \tau$  and  $\tau_1$ , such that, for all  $n \geq 1$ ,

$$\mathbb{Q}[e^{\pm t(\ln W_n - \mathbb{Q}[\ln W_n])}] \leq \begin{cases} \exp(2n\tau_1\beta\rho t^\rho) & \text{if } t > \frac{t_0}{\beta}, \\ \exp(nAt^2) & \text{if } 0 < t \leq \frac{t_0}{\beta}, \end{cases} \quad (6.24)$$

and

$$\mathbb{Q}\left[\pm \frac{1}{n}(\ln W_n - \mathbb{Q}[\ln W_n]) > x\right] \leq \begin{cases} \exp(-nR_1 x^Q) & \text{if } x > 2\rho\beta\tau_1 t_0^{\rho-1}, \\ \exp(-nBx^2) & \text{if } 0 < x \leq 2\rho\beta\tau_1 t_0^{\rho-1}, \end{cases} \quad (6.25)$$

where  $R_1 > 0$  is such that  $\beta(2\rho\tau_1)^{\frac{1}{\rho}}(QR_1)^{\frac{1}{Q}} = 1$ .

If we are not interested in the values of constants, then we have

COROLLARY 6.7 Under the conditions of Theorem 6.6, there exist constants  $c_1, c_2 > 0$ , depending only on  $\beta, K_0, Q$  and  $R$ , such that:

$$\mathbb{Q}[e^{\pm t(\ln W_n - \mathbb{Q}[\ln W_n])}] \leq \begin{cases} \exp(nc_1 t^\rho) & \text{if } t > 1, \\ \exp(nc_1 t^2) & \text{if } 0 < t \leq 1, \end{cases} \quad (6.26)$$

and

$$\mathbb{Q}\left[\pm \frac{1}{n}(\ln W_n - \mathbb{Q}[\ln W_n]) > x\right] \leq \begin{cases} \exp(-nc_2 x^Q) & \text{if } x > 1, \\ \exp(-nc_2 x^2) & \text{if } 0 < x \leq 1. \end{cases} \quad (6.27)$$

In particular, if  $K_0 := \mathbb{Q}[e^{R|\eta|^2}] < +\infty$  for some  $R > 0$ , then for some constants  $c_1, c_2 > 0$  depending only on  $\beta, K_0$  and  $R$ ,

$$\mathbb{Q}[e^{\pm t(\ln W_n - \mathbb{Q}[\ln W_n])}] \leq \exp(nc_1 t^2) \text{ for all } t \in \mathbb{R}, \quad (6.28)$$

and

$$\mathbb{Q}\left[\pm \frac{1}{n}(\ln W_n - \mathbb{Q}[\ln W_n]) > x\right] \leq \exp(-nc_2 x^2) \text{ for all } x > 0. \quad (6.29)$$

REMARK 6.8 If the environment is bounded or gaussian, the inequality (6.28) was proved in [13], Corollary 2.5, as a corollary of a general concentration result.

**Proof of Theorem 6.6.** Let  $\tau_1 > \tau$ . By Lemmas 3.5 and 3.3, writing  $a = K_0(\frac{2}{R})^{\frac{1}{Q-1}}$ , we have

$$\mathbb{Q}[e^{t|\eta|}] \leq 1 + K_0 + at^\rho e^{\tau_1 t^\rho} \leq e^{\tau_1 t^\rho} \forall t \geq t_0,$$

for some  $t_0 = t_0(K_0, \rho, \tau, \tau_1) > 1$ . Hence  $\lambda(\pm t) \leq \tau_1 t^\rho \forall t \geq t_0$ , so by Lemma 6.4,

$$\mathbb{Q}_{j-1}[\exp(\pm t V_{n,j})] \leq \exp(L(\pm t)) \leq \exp(2\tau_1 \beta^\rho t^\rho) \text{ for all } t \geq \frac{t_0}{\beta}. \quad (6.30)$$

We apply Lemma 2.3 with  $I = ]\frac{t_0}{\beta}, +\infty[$ , and with the aid of Lemma 3.4, we conclude that

$$\mathbb{Q}[e^{\pm t(\ln W_n - \mathbb{Q}[\ln W_n])}] \leq \exp(2n\tau_1 \beta^\rho t^\rho) \text{ if } t > \frac{t_0}{\beta}, \quad (6.31)$$

and

$$\mathbb{Q}\left[\pm \frac{1}{n}(\ln W_n - \mathbb{Q}[\ln W_n]) > x\right] \leq \exp(-nR_1 x^Q) \text{ if } x > 2\rho\tau_1 \beta t_0^{\rho-1}. \quad (6.32)$$

Clearly, the condition  $\mathbb{Q}[e^{R|\eta|^Q}] < \infty$  implies  $\mathbb{Q}[e^{\beta|\eta|}] < \infty$ . Let  $K = 2 \exp(\lambda(\beta) + \lambda(-\beta))$ . By Theorem 6.1, (6.4), and Corollary 6.2, (6.7),  $\forall \varepsilon > 0$ ,

$$\mathbb{Q}[e^{\pm t(\ln W_n - \mathbb{Q}[\ln W_n])}] \leq \exp\left(\frac{nKt^2}{1-t}\right) \leq \exp\left(\frac{nKt^2}{1-\varepsilon}\right) \text{ if } 0 < t \leq \varepsilon, \quad (6.33)$$

and

$$\mathbb{Q}\left[\pm \frac{1}{n}(\ln W_n - \mathbb{Q}[\ln W_n]) > x\right] \leq \exp\left(-\frac{nx^2}{4K(1+\varepsilon)}\right) \text{ if } 0 < x < \delta(K, \varepsilon), \quad (6.34)$$

for some  $\delta(K, \varepsilon)$  small enough. In the following, we take  $\varepsilon = \frac{1}{2}$  and  $\delta = \delta(K, \frac{1}{2})$ . If  $\frac{1}{2} \leq t \leq \frac{t_0}{\beta}$ , then

$$\mathbb{Q}[e^{\pm t(\ln W_n - \mathbb{Q}[\ln W_n])}] \leq \mathbb{Q}\left[e^{\frac{t_0}{\beta}|\ln W_n - \mathbb{Q}[\ln W_n]|}\right] \leq 2 \exp(2n\tau_1 t_0^\rho) \leq \exp(n(4\ln 2 + 8\tau_1 t_0^\rho)t^2). \quad (6.35)$$

Combining (6.31), (6.33) and (6.35) gives (6.24), with  $A = \max(4K, 4\ln 2 + 8\tau_1 t_0^\rho)$ .

If  $\delta \leq x \leq x_0 := 2\rho\tau_1 \beta t_0^{\rho-1}$ , then by (6.5),

$$\mathbb{Q}\left[\pm \frac{1}{n}(\ln W_n - \mathbb{Q}[\ln W_n]) > x\right] \leq \exp\left(-n(\sqrt{\delta+K} - \sqrt{K})^2\right) \leq \exp\left(-n\frac{(\sqrt{\delta+K} - \sqrt{K})^2 x^2}{x_0^2}\right). \quad (6.36)$$

Combining (6.32), (6.34) and (6.36) gives (6.25), with  $B = \min\left(\frac{1}{6K}, \frac{(\sqrt{\delta+K} - \sqrt{K})^2}{x_0^2}\right)$ . ■

## 7 Free energy of directed polymers: convergence rates

It is well known that the sequence  $\mathbb{Q}[\ln W_n(\beta)]$  is superadditive, hence the limit

$$p_-(\beta) = \lim_{n \rightarrow \infty} \frac{1}{n} \mathbb{Q}[\ln(W_n(\beta))] = \sup_n \frac{1}{n} \mathbb{Q}[\ln(W_n(\beta))] \in [-\infty, 0] \quad (7.1)$$

exists<sup>2</sup>. As an immediate consequence of (7.1) and (6.20), we have:

LEMMA 7.1 *Assume that  $\mathbb{Q}[e^{\beta|\eta|}] < +\infty$ . Then*

$$p_-(\beta) = \lim_{n \rightarrow \infty} \frac{1}{n} \ln(W_n(\beta)) \in [\beta \mathbb{Q}[\eta] - \lambda(\beta), 0] \quad Q\text{-a.s. and in } L^p, \quad \forall p \geq 1. \quad (7.2)$$

The inequality  $p_-(\beta) \leq 0$  was already indicated in (7.1); it follows from the fact that

$$\mathbb{Q}[\ln(W_n(\beta))] \leq \ln \mathbb{Q}[W_n(\beta)] = 0.$$

The inequality  $p_-(\beta) \geq \beta \mathbb{Q}[\eta] - \lambda(\beta)$  also comes directly from the definition, as

$$\mathbb{Q}[\ln(W_n)] \geq \mathbb{Q}P[\beta H_n - n\lambda(\beta)] = P\mathbb{Q}[\beta H_n - n\lambda(\beta)] = n(\beta \mathbb{Q}[\eta] - \lambda(\beta)).$$

The a.s. convergence was proved in [11], under the stronger condition that  $\mathbb{Q}[e^{\beta|\eta|}] < +\infty$  for all  $\beta > 0$ ; actually their proof is valid under the condition that  $\mathbb{Q}[e^{3\beta|\eta|}] < +\infty$ . We shall give an estimation of the rate of convergence, for each of the convergences in probability, a.s., and in  $L^p$  ( $p \geq 1$ ): cf. Theorems 7.2 and 7.5.

We first consider the rate of convergence in probability. Recall that the condition  $\mathbb{Q}[e^{\beta|\eta|}] < +\infty$  is equivalent to  $\lambda(\pm\beta) < \infty$ .

**THEOREM 7.2** *If  $K := 2 \exp(\lambda(-\beta)) + \lambda(\beta) < \infty$ , then  $\forall \delta \in ]0, 1[$ ,  $\forall x > 0$ , there exists  $n_0 = n_0(\delta, x) > 0$  such that  $\forall n \geq n_0$ ,*

$$\mathbb{Q}\left[\left|\frac{1}{n} \ln W_n - p_-(\beta)\right| > x\right] \leq 2 \exp\left(-n \left(\sqrt{(1-\delta)x+K} - \sqrt{K}\right)^2\right). \quad (7.3)$$

Consequently,

$$\mathbb{Q}\left[\left|\frac{1}{n} \ln W_n - p_-(\beta)\right| > x\right] \leq \begin{cases} 2 \exp\left(-\frac{n(1-\delta)^2 x^2}{K(1+\sqrt{2})^2}\right) & \text{if } x(1-\delta) \leq K, \\ 2 \exp\left(-\frac{n(1-\delta)x}{(1+\sqrt{2})^2}\right) & \text{if } x(1-\delta) > K. \end{cases} \quad (7.4)$$

In particular (take  $\delta = \frac{1}{2}$ ),  $\forall x \in ]0, 2K]$ , there exists  $n_0 = n_0(x) > 0$  such that  $\forall n \geq n_0$ ,

$$\mathbb{Q}\left[\left|\frac{1}{n} \ln W_n - p_-(\beta)\right| > x\right] \leq 2 \exp\left(-\frac{nx^2}{4K(1+\sqrt{2})^2}\right). \quad (7.5)$$

**Proof.** Let  $\delta \in ]0, 1[$ , and  $x > 0$ . Let  $n_0 = n_0(\delta, x)$  be large enough such that for any  $n \geq n_0$ ,

$$0 \leq p_-(\beta) - \frac{1}{n} \mathbb{Q}[\ln(W_n(\beta))] < \delta x.$$

Then  $\forall n \geq n_0$ ,

$$\mathbb{Q}\left[\left|\frac{1}{n} \ln W_n - p_-(\beta)\right| > x\right] \leq \mathbb{Q}\left[\left|\frac{1}{n} \ln W_n - \frac{1}{n} \mathbb{Q}[\ln W_n]\right| > (1-\delta)x\right].$$

Therefore the conclusion follows from Theorem 6.1. ■

---

<sup>2</sup>In the literature,  $p(\beta)$  is often used to denote the limit of the un-normalized free energy:  $p(\beta) = \lim_{n \rightarrow \infty} \frac{1}{n} \mathbb{Q}[\ln(Z_n(\beta))]$ . We use the symbol  $p_-(\beta)$  to indicate that  $p_-(\beta) \leq 0$ . Of course  $p_-(\beta) = p(\beta) - \lambda(\beta)$ .

We next consider the rate of convergence in mean. To this end, we first introduce some notations. We note  $P^x$  the law of the simple random walk on  $\mathbb{Z}^d$  starting at  $x$ , and  $L_m = \{x \in \mathbb{Z}^d, P(\omega_m = x) > 0\}$ . In addition to the partition function  $W_n$ , we define the partition function starting from  $x$ :

$$W_n(x) = W_n(x; \eta) = P^x \left[ \exp \left( \beta \sum_{j=1}^n \eta(j, \omega_j) - n\lambda(\beta) \right) \right], \quad (7.6)$$

and the point to point partition function

$$W_n(x, y) = W_n(x, y; \eta) = P^x \left[ \exp \left( \beta \sum_{j=1}^n \eta(j, \omega_j) - n\lambda(\beta) \right) \mathbf{1}_{\omega_n=y} \right]. \quad (7.7)$$

Let  $\tau_n$  be the time shift of ordre  $n$  on the environment:

$$(\tau_n \eta)(k, x) = \eta(k + n, x) \quad (x \in \mathbb{Z}^d, k \geq 1).$$

Then we have

$$W_{n+k} = \sum_{x \in L_n} W_n(0, x; \eta) W_k(x; \tau_n \eta). \quad (7.8)$$

**LEMMA 7.3 (Rate of convergence in mean)** *If  $K := 2 \exp(\lambda(-\beta)) + \lambda(\beta) < \infty$ , then for each  $n \in \mathbb{N}^*$ ,*

$$0 \leq p_-(\beta) - \frac{1}{n} \mathbb{Q}[\ln(W_n(\beta))] \leq 2\sqrt{K} \sqrt{\frac{d \ln(2n)}{n}} + \frac{d \ln(2n)}{n}. \quad (7.9)$$

**Proof.** We adapt the proof of Proposition 2.4 of [7]. Let  $\varepsilon \in ]0, 1[$ . Using (7.8) and the subadditivity of the function  $u \mapsto u^\varepsilon$ , we get

$$W_{n+k}^\varepsilon \leq \sum_{x \in L_n} W_n^\varepsilon(0, x; \eta) W_k^\varepsilon(x; \tau_n \eta).$$

Integrating with respect to the environment we have

$$\mathbb{Q}[W_{n+k}^\varepsilon] \leq |L_n| \mathbb{Q}[W_n^\varepsilon] \mathbb{Q}[W_k^\varepsilon] \leq (2n)^d \mathbb{Q}[W_n^\varepsilon] \mathbb{Q}[W_k^\varepsilon].$$

Therefore  $h_\varepsilon(n) := \ln \mathbb{Q}[W_n^\varepsilon]$  ( $\geq \varepsilon \mathbb{Q}[\ln W_n]$ ) satisfies

$$h_\varepsilon(n+k) \leq h_\varepsilon(n) + h_\varepsilon(k) + d \ln(2n) \quad \forall n, k \geq 1.$$

Set  $h(\varepsilon) = \limsup_{n \rightarrow +\infty} \frac{h_\varepsilon(n)}{n}$  (in fact by Hammersley's (1962, [20]) theorem on sub-additive functions, the limit exists, although we shall not use it). By the preceding recurrence relation, we have

$$h_\varepsilon(nm) \leq mh_\varepsilon(n) + (m-1)d \ln(2n), \quad n, m \geq 1.$$

Dividing this inequality by  $nm$  and letting  $m \rightarrow \infty$ , we see that

$$h(\varepsilon) \leq \frac{h_\varepsilon(n) + d \ln(2n)}{n}, \quad \forall n \geq 1.$$

By Theorem 6.1,

$$h_\varepsilon(n) = \ln \mathbb{Q}[\exp(\varepsilon(\ln W_n - \mathbb{Q}[\ln W_n]))] + \varepsilon \mathbb{Q}[\ln W_n] \leq \frac{nK\varepsilon^2}{1-\varepsilon} + \varepsilon \mathbb{Q}[\ln W_n].$$

As  $p_-(\beta) \leq \frac{h(\varepsilon)}{\varepsilon}$ , it follows that

$$p_-(\beta) \leq \frac{K\varepsilon}{1-\varepsilon} + \frac{\mathbb{Q}[\ln W_n]}{n} + \frac{d \ln(2n)}{n\varepsilon} \quad \forall \varepsilon \in ]0, 1[. \quad (7.10)$$

Let  $g(\varepsilon) = \frac{K\varepsilon}{1-\varepsilon} + \frac{d_n}{\varepsilon}$ , where  $d_n = \frac{d \ln(2n)}{n}$ ,  $\varepsilon \in ]0, 1[$ . Then  $g'(\varepsilon) = \frac{K}{(1-\varepsilon)^2} - \frac{d_n}{\varepsilon^2} = 0$  if and only if  $\varepsilon = \frac{\sqrt{d_n}}{\sqrt{K} + \sqrt{d_n}}$ . For  $\varepsilon = \frac{\sqrt{d_n}}{\sqrt{K} + \sqrt{d_n}}$ ,  $g(\varepsilon) = 2\sqrt{Kd_n} + d_n$ ; therefore taking  $\varepsilon = \frac{\sqrt{d_n}}{\sqrt{K} + \sqrt{d_n}}$  in (7.10), we obtain

$$p_-(\beta) \leq 2\sqrt{Kd_n} + d_n + \frac{\mathbb{Q}[\ln W_n]}{n},$$

that is,

$$p_-(\beta) \leq 2\sqrt{K} \sqrt{\frac{d \ln(2n)}{n}} + \frac{d \ln(2n)}{n} + \frac{\mathbb{Q}[\ln W_n]}{n}.$$

■

As an immediate consequence of the preceding lemma, we have:

COROLLARY 7.4 *If  $K := 2 \exp(\lambda(-\beta)) + \lambda(\beta)) < \infty$ , then*

$$\limsup_{n \rightarrow +\infty} \sqrt{\frac{n}{\ln n}} \left| \frac{\mathbb{Q}[\ln W_n]}{n} - p_-(\beta) \right| \leq 2\sqrt{Kd}. \quad (7.11)$$

We finally consider the rate of convergence, with probability 1 and in  $L^p$ . As usual,  $\|\cdot\|_p$  denotes the  $L^p$  norm.

THEOREM 7.5 *(Rate of convergence, a.s. and in  $L^p$ ) If  $K := 2 \exp(\lambda(-\beta)) + \lambda(\beta)) < \infty$ , then*

$$\limsup_{n \rightarrow +\infty} \sqrt{\frac{n}{\ln n}} \left| \frac{\ln W_n}{n} - p_-(\beta) \right| \leq 2\sqrt{K}(1 + \sqrt{d}) \text{ a.s.,} \quad (7.12)$$

and

$$\limsup_{n \rightarrow +\infty} \sqrt{\frac{n}{\ln n}} \left\| \frac{\ln W_n}{n} - p_-(\beta) \right\|_p \leq 2\sqrt{Kd}, \quad \forall p \geq 1. \quad (7.13)$$

**Proof.** We write

$$\frac{\ln W_n}{n} - p_-(\beta) = \left( \frac{\ln W_n}{n} - \frac{\mathbb{Q}[\ln W_n]}{n} \right) + \left( \frac{\mathbb{Q}[\ln W_n]}{n} - p_-(\beta) \right).$$

Then combining (6.21) of Theorem 6.5 and (7.11) of Corollary 7.4, we get (7.12). Again by Theorem 6.5, we know that for every  $p \geq 1$ ,

$$\left\| \frac{\ln W_n - \mathbb{Q}[\ln W_n]}{n} \right\|_p = O(n^{-1/2}) = o\left(\sqrt{\frac{\ln n}{n}}\right),$$

so that (7.13) is a consequence of Corollary 7.4. ■

REMARK 7.6 Carmona and Hu have proved in [7] that if the environment is gaussian, then for any  $\varepsilon > 0$ ,

$$\left| \frac{\ln W_n}{n} - p_-(\beta) \right| \leq n^{-(\frac{1}{2}-\varepsilon)} \text{ for } n \text{ big enough.}$$

Our estimation is sharper since  $n^{(\frac{1}{2}-\varepsilon)}$  is replaced by  $\sqrt{\frac{\ln n}{n}}$ .

## 8 Expression of the free energy by multiplicative cascades

In this section we shall prove that the free energy  $p_-(\beta)$  can be expressed in terms of the free energies of some generalized multiplicative cascades. The expression is interesting because we know more information on the free energies of multiplicative cascades. The model of multiplicative cascades was first introduced by Mandelbrot (1974, [30]); it has been well studied in the literature: see for example Kahane and Peyrière (1976, [25]), Durrett and Liggett (1981, [17]), Guivarc'h (1990, [19]), Franchi (1993, [18]); for a generalized version and closely related topics, see Liu (2000, [29]).

In [13], Comets and Vargas introduced a generalized multiplicative cascade (cf. [29])  $(W_{m,n}^{\text{tree}})_{n \geq 1}$  associated to the random vector  $(W_m(0, x))_{x \in L_m}$ , where we recall that

$$W_m(0, x) = P[\exp(\beta H_m(\omega) - m\lambda(\beta)); \omega_m = x]. \quad (8.1)$$

The associated free energy is

$$p_m^{tree}(\beta) = \inf_{\theta \in ]0,1]} v_m(\theta), \text{ with } v_m(\theta) = \frac{1}{\theta} \ln \left( \mathbb{Q} \left[ \sum_{x \in L_m} W_m(0, x)^\theta \right] \right). \quad (8.2)$$

Comets et Vargas proved that

$$p_-(\beta) \leq \inf_{m \geq 1} \frac{1}{m} p_m^{tree}(\beta) = \lim_{m \rightarrow +\infty} \frac{1}{m} p_m^{tree}(\beta), \quad (8.3)$$

and that the equality holds if the environment is gaussian or bounded. Here we prove that the equality holds for general environment.

**THEOREM 8.1** *Assume that  $\mathbb{Q}[e^{\beta|\eta|}] < +\infty$ . Then*

$$p_-(\beta) = \inf_{m \geq 1} \frac{1}{m} p_m^{tree}(\beta). \quad (8.4)$$

**Proof.** For the sake of completeness, we recall the argument of Comets-Vargas for the inequality (8.3). Using the point to point partition functions defined by (7.7), we have

$$W_{mn} = \sum_{x_1, \dots, x_n \in \mathbb{Z}^d} W_m(0, x_1) W_m(x_1, x_2; \tau_m \eta) \cdots W_m(x_{n-1}, x_n; \tau_{(n-1)m} \eta). \quad (8.5)$$

Let  $\theta \in ]0, 1[$  and  $m \in \mathbb{N}^*$ . By the subadditivity of the function  $u \mapsto u^\theta$  and Jensen's inequality, we obtain:

$$\begin{aligned} \frac{1}{nm} \mathbb{Q}[\ln W_{nm}] &= \frac{1}{m\theta n} \mathbb{Q}[\ln W_{nm}^\theta] \\ &\leq \frac{1}{m\theta n} \mathbb{Q} \left[ \ln \sum_{x_1, \dots, x_n \in \mathbb{Z}^d} W_m^\theta(0, x_1) W_m^\theta(x_1, x_2; \tau_m \eta) \cdots W_m^\theta(x_{n-1}, x_n; \tau_{(n-1)m} \eta) \right] \\ &\leq \frac{1}{m\theta n} \ln \mathbb{Q} \left[ \sum_{x_1, \dots, x_n \in \mathbb{Z}^d} W_m^\theta(0, x_1) W_m^\theta(x_1, x_2; \tau_m \eta) \cdots W_m^\theta(x_{n-1}, x_n; \tau_{(n-1)m} \eta) \right]. \end{aligned}$$

By induction on  $n$  it is easy to see that

$$\mathbb{Q} \left[ \sum_{x_1, \dots, x_n \in \mathbb{Z}^d} W_m^\theta(0, x_1) W_m^\theta(x_1, x_2; \tau_m \eta) \cdots W_m^\theta(x_{n-1}, x_n; \tau_{(n-1)m} \eta) \right] = \left( \mathbb{Q} \left[ \sum_{x \in L_m} W_m^\theta(0, x) \right] \right)^n,$$

therefore

$$\frac{1}{nm} \mathbb{Q}[\ln W_{nm}] \leq \frac{1}{m\theta} \ln \left( \mathbb{Q} \left[ \sum_{x \in L_m} W_m^\theta(0, x) \right] \right).$$

Letting  $n \rightarrow \infty$  gives

$$p_-(\beta) \leq \frac{1}{m\theta} \ln \left( \mathbb{Q} \left[ \sum_{x \in L_m} W_m^\theta(0, x) \right] \right);$$

then taking the infimum over all  $\theta \in ]0, 1]$  gives

$$p_-(\beta) \leq \frac{1}{m} p_m^{tree}(\beta).$$

Now we prove the reverse inequality. As  $W_m(0, x) \leq W_m$  for every  $x$ , we have, for  $\theta \in ]0, 1[$ ,

$$v_m(\theta) \leq \frac{1}{\theta} \ln (\mathbb{Q}[|L_m| W_m^\theta]),$$

where  $|L_m|$  is the cardinality of  $L_m$ . Writing

$$\mathbb{Q}[W_m^\theta] = e^{\theta \mathbb{Q}[\ln W_m]} \mathbb{Q}[\exp(\theta(\ln W_m - \mathbb{Q}[\ln W_m]))],$$

we get

$$v_m(\theta) \leq \frac{1}{\theta} \ln |L_m| + \mathbb{Q}[\ln W_m] + \frac{1}{\theta} \ln (\mathbb{Q}[\exp(\theta(\ln W_m - \mathbb{Q}[\ln W_m]))]).$$

Recall that by Theorem 6.1, for every  $\theta \in ]0, 1[$ ,

$$\mathbb{Q}[\exp(\theta(\ln W_m - \mathbb{Q}[\ln W_m]))] \leq e^{\frac{mK\theta^2}{1-\theta}},$$

with  $K = 2 \exp(\lambda(-\beta)) + \lambda(\beta)$ . Therefore for any  $m \geq 1$ ,

$$\inf_{m \geq 1} \frac{1}{m} p_m^{tree}(\beta) \leq \frac{1}{m} p_m^{tree}(\beta) \leq \frac{1}{m} v_m(\theta) \leq \frac{1}{m\theta} \ln |L_m| + \frac{1}{m} \mathbb{Q}[\ln W_m] + \frac{K\theta}{1-\theta}.$$

Letting  $m \rightarrow \infty$  and using the fact that  $|L_m| \leq (2m)^d$ , we obtain that

$$\inf_{m \geq 1} \frac{1}{m} p_m^{tree}(\beta) \leq p_-(\beta) + \frac{K\theta}{1-\theta}.$$

This gives the desired result as  $\theta \in ]0, 1[$  is arbitrary. ■

## References

- [1] Kazuoki Azuma. Weighted sums of certain dependent random variables. *Tôhoku Math. J. (2)*, 19:357–367, 1967.
- [2] Vidmantas Bentkus. On Hoeffding’s inequalities. *Ann. Probab.*, 32(2):1650–1673, 2004.
- [3] S. Bernstein. Sur une modification de l’inégalité de Tchebichef. *Annals Science Institute Sav. Ukraine, Sect. Math.*, I, 1924.
- [4] Erwin Bolthausen. A note on the diffusion of directed polymers in a random environment. *Comm. Math. Phys.*, 123(4):529–534, 1989.
- [5] Philippe Carmona, Francesco Guerra, Yueyun Hu, and Olivier Menjane. Strong disorder for a certain class of directed polymers in a random environment. *J. Theoret. Probab.*, 19(1):134–151, 2006.
- [6] Philippe Carmona and Yueyun Hu. On the partition function of a directed polymer in a Gaussian random environment. *Probab. Theory Related Fields*, 124(3):431–457, 2002.
- [7] Philippe Carmona and Yueyun Hu. Fluctuation exponents and large deviations for directed polymers in a random environment. *Stochastic Process. Appl.*, 112(2):285–308, 2004.
- [8] Herman Chernoff. A note on an inequality involving the normal distribution. *Ann. Probab.*, 9(3):533–535, 1981.
- [9] Yuan Shih Chow and Henry Teicher. *Probability theory*. Springer Texts in Statistics. Springer-Verlag, New York, third edition, 1997. Independence, interchangeability, martingales.
- [10] Fan Chung and Linyuan Lu. Concentration inequalities and martingale inequalities: a survey. *Internet Math.*, 3(1):79–127, 2006.
- [11] Francis Comets, Tokuzo Shiga, and Nobuo Yoshida. Directed polymers in a random environment: path localization and strong disorder. *Bernoulli*, 9(4):705–723, 2003.
- [12] Francis Comets, Tokuzo Shiga, and Nobuo Yoshida. Probabilistic analysis of directed polymers in a random environment: a review. In *Stochastic analysis on large scale interacting systems*, volume 39 of *Adv. Stud. Pure Math.*, pages 115–142. Math. Soc. Japan, Tokyo, 2004.
- [13] Francis Comets and Vincent Vargas. Majorizing multiplicative cascades for directed polymers in random media. *ALEA Lat. Am. J. Probab. Math. Stat.*, 2:267–277 (electronic), 2006.
- [14] Francis Comets and Nobuo Yoshida. Directed polymers in random environment are diffusive at weak disorder. *Ann. Probab.*, 34(5):1746–1770, 2006.

- [15] H. Cramér. Sur un nouveau théorème-limite de la théorie des probabilités. *Actualités Sci. Indust. Hermann, Paris*, 736, 5C23, 1938.
- [16] Victor H. de la Peña. A general class of exponential inequalities for martingales and ratios. *Ann. Probab.*, 27(1):537–564, 1999.
- [17] Richard Durrett and Thomas M. Liggett. Fixed points of the smoothing transformation. *Z. Wahrsch. Verw. Gebiete*, 64(3):275–301, 1983.
- [18] J. Franchi. Chaos multiplicatif: un traitement simple et complet de la fonction de partition. In *Séminaire de Probabilités, XXIX*, volume 1613 of *Lecture Notes in Math.*, pages 194–201. Springer, Berlin, 1995.
- [19] Yves Guivarc'h. Sur une extension de la notion de loi semi-stable. *Ann. Inst. H. Poincaré Probab. Statist.*, 26(2):261–285, 1990.
- [20] J. M. Hammersley. Generalization of the fundamental theorem on sub-additive functions. *Proc. Cambridge Philos. Soc.*, 58:235–238, 1962.
- [21] T. E. Harris. Branching processes. *Ann. Math. Statistics*, 19:474–494, 1948.
- [22] Wassily Hoeffding. Probability inequalities for sums of bounded random variables. *J. Amer. Statist. Assoc.*, 58:13–30, 1963.
- [23] Henley C. L. Huse, D. A. Pinning and roughening of domain walls in ising systems due to random impurities. *Phys. Rev. Lett.*, 54:2708–2711, 1985.
- [24] J. Z. Imbrie and T. Spencer. Diffusion of directed polymers in a random environment. *J. Statist. Phys.*, 52(3-4):609–626, 1988.
- [25] J.-P. Kahane and J. Peyrière. Sur certaines martingales de Benoit Mandelbrot. *Advances in Math.*, 22(2):131–145, 1976.
- [26] Michel Ledoux. Concentration of measure and logarithmic Sobolev inequalities. In *Séminaire de Probabilités, XXXIII*, volume 1709 of *Lecture Notes in Math.*, pages 120–216. Springer, Berlin, 1999.
- [27] Emmanuel Lesigne and Dalibor Volný. Large deviations for martingales. *Stochastic Process. Appl.*, 96(1):143–159, 2001.
- [28] Quansheng Liu. The growth of an entire characteristic function and the tail probabilities of the limit of a tree martingale. In *Trees (Versailles, 1995)*, volume 40 of *Progr. Probab.*, pages 51–80. Birkhäuser, Basel, 1996.
- [29] Quansheng Liu. On generalized multiplicative cascades. *Stochastic Process. Appl.*, 86(2):263–286, 2000.
- [30] Benoit Mandelbrot. Multiplications aléatoires itérées et distributions invariantes par moyenne pondérée aléatoire: quelques extensions. *C. R. Acad. Sci. Paris Sér. A*, 278:289–292 and 355–358, 1974.
- [31] Valentin V. Petrov. *Limit theorems of probability theory*, volume 4 of *Oxford Studies in Probability*. The Clarendon Press Oxford University Press, New York, 1995. Sequences of independent random variables, Oxford Science Publications.
- [32] B. Ramachandran. On the order and the type of entire characteristic functions. *Ann. Math. Statist.*, 33:1238–1255, 1962.
- [33] Feng-Yu Wang. *Functional inequalities, Markov semigroups and spectral theory*. Science Press, Beijing, New York, 2005.